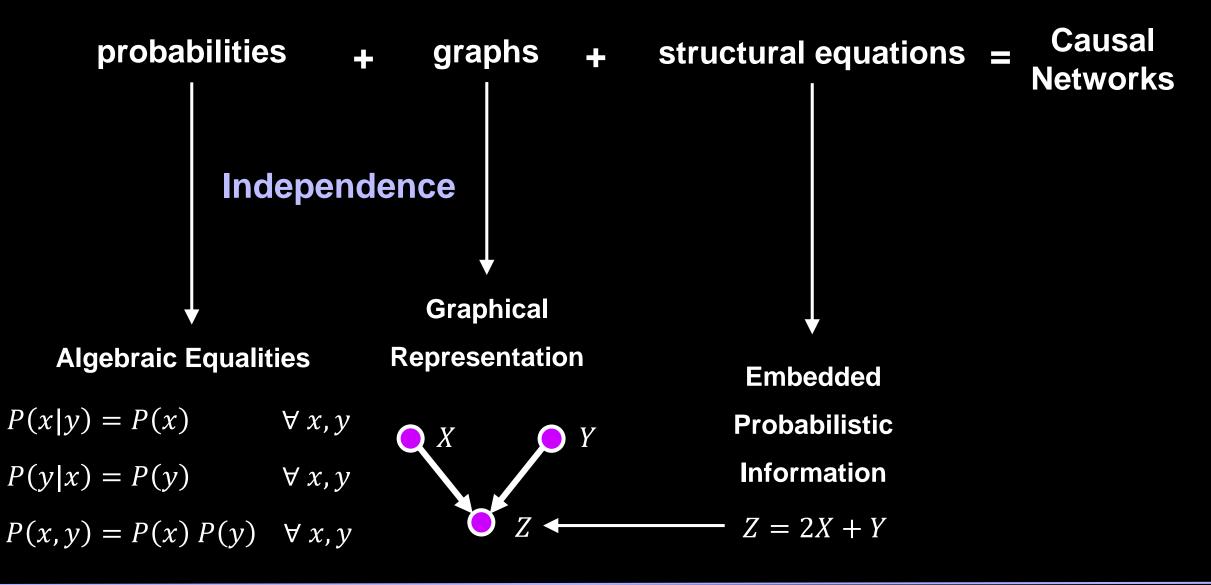


2 GRAPHICAL MODELS AND THEIR APPLICATIONS

2.1 CONNECTING MODELS TO DATA

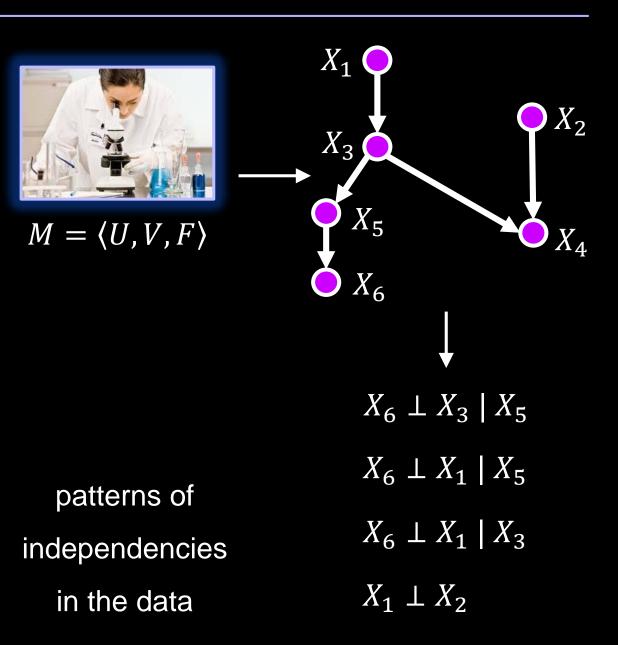
In **Part 1**, we introduced



2.1 CONNECTING MODELS TO DATA

The researcher who has scientific knowledge in the form of **structural equation model** is able to predict patterns of independencies in the data, based solely on the structure of the model's graph, without relying on any quantitative information carried by the equations or by the distributions of errors.

X 1	X ₂	Х 3	X 4	X ₅	Х ₆
0	1	1	1	0	1
0	1	0	1	1	1
1	0	1	1	0	0
0	0	1	0	1	1
0	1	0	0	0	0
1	1	0	1	0	1
1	0	1	1	1	1
0	1	1	1	0	1



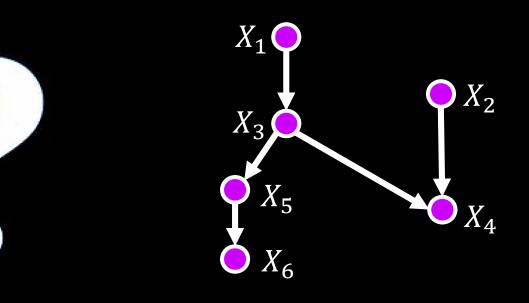
2.1 CONNECTING MODELS TO DATA



X 1	X ₂	X ₃	X 4	X ₅	X ₆
0	1	1	1	0	1
0	1	0	1	1	1
1	0	1	1	0	0
0	0	1	0	1	1
0	1	0	0	0	0
1	1	0	1	0	1
1	0	1	1	1	1
0	1	1	1	0	1

Data Set

Conversely, it means that observing **patterns of independencies** in the data enables us to say something about whether a hypothesized model is correct.



Hypothesized Model

represent the causal story behind the data X ₂ *X*₄ X ₅ *X*₆ *X*₁ X ₃ $M = \langle U, V, F \rangle$ **Causal Model** Data Set mechanism by which data

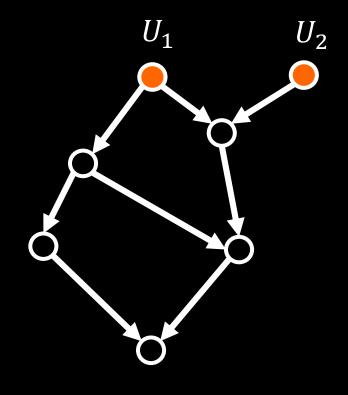
were generated

Given a truly complete causal model for, say, math test score in high school juniors, and given complete list of values for every exogenous variable in that model, we could theoretically generate a data point (i.e., a test score), for each individual (student).

 U_1 U_2 $M = \langle U, V, F \rangle$ X_2 X_1 $U = \{U_1, U_2\}$ X_4 X_3 $V = \{X_1, X_2, X_3, X_4, X_5\}$ $F = \{f_1, f_2, f_3, f_4, f_5\}$ X_5

Given a truly complete causal model for, say, math test score in high school juniors, and given complete list of values for every exogenous variable in that model, we could theoretically generate a data point (i.e., a test score), for each individual (student).

 $M = \langle U, V, F \rangle$ $U = \{U_1, U_2\}$ $V = \{X_1, X_2, X_3, X_4, X_5\}$ $F = \{f_1, f_2, f_3, f_4, f_5\}$



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$$M = \langle U, V, F \rangle$$

$$U = \{U_1, U_2\}$$

$$V = \{X_1, X_2, X_3, X_4, X_5\}$$

$$F = \{f_1, f_2, f_3, f_4, f_5\}$$

Given a truly complete causal model for, say, math test score in high school juniors, and given complete list of values for every exogenous variable in that model, we could theoretically generate a data point (i.e., a test score), for each individual (student).

$$M = \langle U, V, F \rangle$$

$$U = \{U_1, U_2\}$$

$$X_4 = f_4(X_2)$$

$$V = \{X_1, X_2, X_3, X_4, X_5\}$$

$$F = \{f_1, f_2, f_3, f_4, f_5\}$$

Given a truly complete causal model for, say, math test score in high school juniors, and given complete list of values for every exogenous variable in that model, we could theoretically generate a data point (i.e., a test score), for each individual (student).

 U_1 U_{2} $M = \langle U, V, F \rangle$ X_1 X_2 $U = \{U_1, U_2\}$ X_4 X_3 $V = \{X_1, X_2, X_3, X_4, X_5\}$ $F = \{f_1, f_2, f_3, f_4, f_5\}$ $X_5 = f_5(X_3, X_4)$

Given a truly complete causal model for, say, math test score in high school juniors, and given complete list of values for every exogenous variable in that model, we could theoretically generate a data point (i.e., a test score), for each individual (student).

 U_1 U_2 $M = \langle U, V, F \rangle$ X_1 X_2 $U = \{U_1, U_2\}$ X_4 X_3 $V = \{X_1, X_2, X_3, X_4, X_5\}$ $F = \{f_1, f_2, f_3, f_4, f_5\}$ X_5

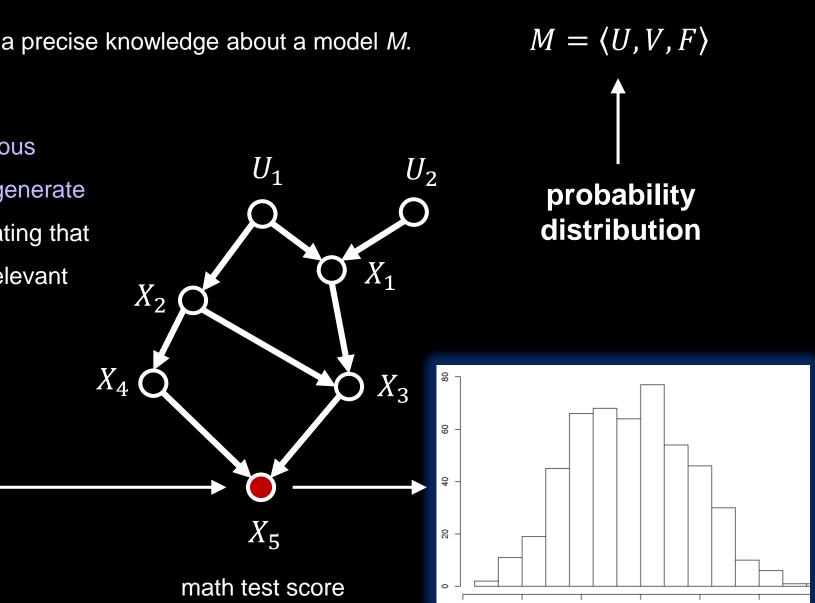
We can compute test score for each student. This would necessitate specifying all factors that may have an effect on a

student's test score, an unrealistic task.

In many cases, we will not have such a precise knowledge about a model M.

We might instead have a probability distribution characterizing the exogenous variables U, which would allow us to generate a distribution of test scores approximating that of the entire student population and relevant subgroups of students.

For each student we do not get the corresponding math test score X_5 but we get a distribution for its value



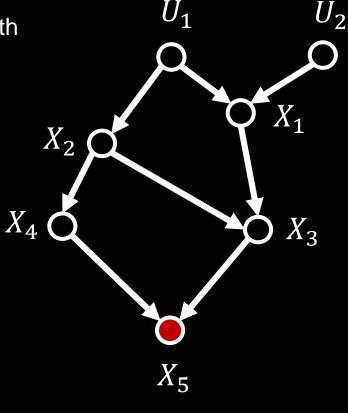
Suppose, however, that we do not have even a probabilistically specified causal model, but only a graphical structure of the model.

We know which variables are caused by which other variables, but we do not know the strength or nature of the relationships.

$$M = \langle U, V, F \rangle \qquad F = ???$$

Unspecified Causal Model

Even with such limited information, we can discern a great deal about the data set generated by the model.

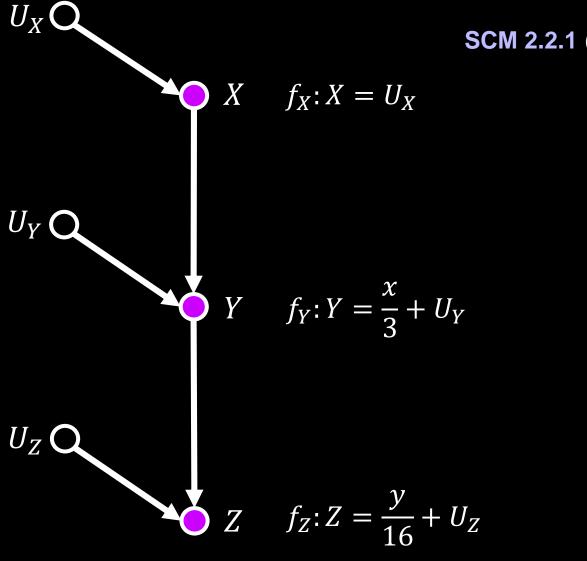


math test score

We can learn which variables in the data set are independent of each other and which are independent of each other conditional on other variables.

These independencies will be true of every data set generated by a causal model with that graphical structure, regardless of the specific functions attached to the SCM.

Consider the following three hypothetical SCMs, all share the same graphical model.



SCM 2.2.1 (School Funding, SAT Scores, and College Acceptance)

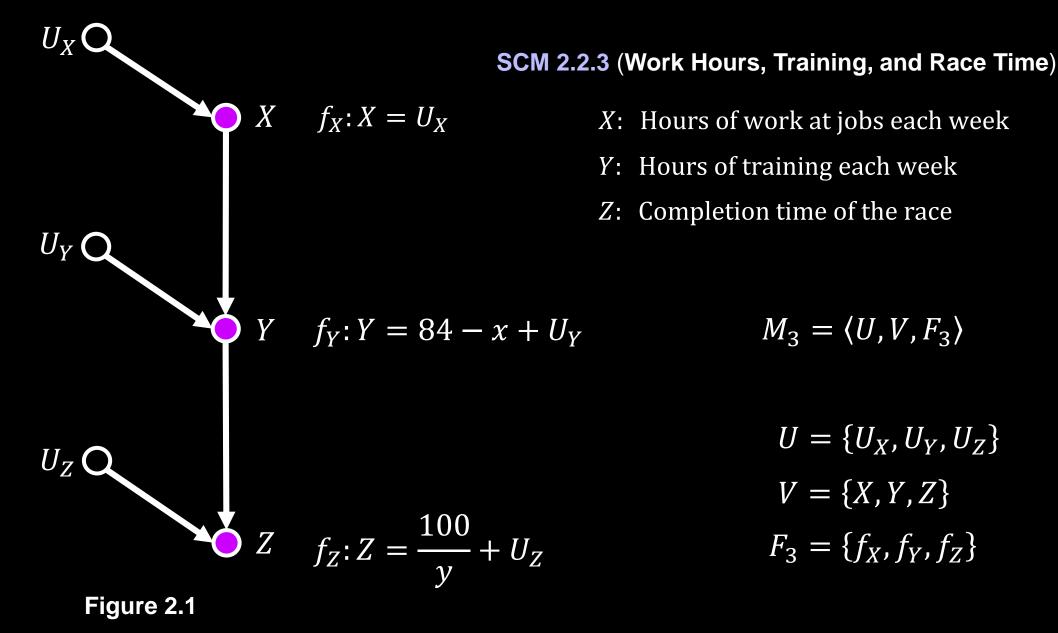
- *X*: High School's funding in dollars
- Y: Average SAT Score
- *Z*: College acceptance rate



 $U = \{U_X, U_Y, U_Z\}$ $V = \{X, Y, Z\}$ $F_1 = \{f_X, f_Y, f_Z\}$

Figure 2.1

 $U_X \mathbf{Q}$ $M_2 = \langle U, V, F_2 \rangle$ SCM 2.2.2 (Switch, Circuit, and Light Bulb) $U = \{U_X, U_Y, U_Z\}$ $f_X: X = U_X$ X: State of a light switch X $V = \overline{\{X, Y, Z\}}$ *Y*: State of the associated electrical circuit *Z*: State of a light bulb $F_2 = \{f_X, f_Y, f_Z\}$ $Y \quad f_Y: Y = \begin{cases} closed & \text{IF} (X = up \text{ AND } U_Y = 0) \text{ OR } (X = down \text{ AND } U_Y = 1) \\ open & \text{otherwise} \end{cases}$ $U_X = \{up, down\} \qquad X = \{up, down\}$ $U_Y = \{0,1\} \qquad Y = \{closed, open\}$ $U_Z C$ $U_Z = \{0,1\} \qquad \qquad \overline{Z} = \{on, off\}$ $\sum_{Z \in Z} f_Z : Z = \begin{cases} on & \text{IF} (Y = closed \text{ AND } U_Z = 0) \text{ OR } (Y = open \text{ AND } U_Z = 1) \\ of f \text{ otherwise} \end{cases}$



$$U_{X} \qquad M_{1} = \langle U, V, F_{1} \rangle \qquad U = \{U_{X}, U_{Y}, U_{Z} \}$$

$$M_{2} = \langle U, V, F_{2} \rangle \qquad U = \{U_{X}, U_{Y}, U_{Z} \}$$

$$M_{3} = \langle U, V, F_{3} \rangle \qquad V = \{X, Y, Z \}$$

$$M_{3} = \langle U, V, F_{3} \rangle \qquad V = \{X, Y, Z \}$$

$$M_{1} = \langle U, V, F_{1} \rangle \qquad U = \{U_{X}, U_{Y}, U_{Z} \}$$

$$V = \{X, Y, Z \}$$

$$M_{1} = \langle U, V, F_{1} \rangle \qquad V = \{U_{X}, U_{Y}, U_{Z} \}$$

$$V = \{X, Y, Z \}$$

$$M_{1} = \langle U, V, F_{1} \rangle \qquad V = \{U_{X}, U_{Y}, U_{Z} \}$$

$$V = \{X, Y, Z \}$$

$$M_{1} = \langle U, V, F_{1} \rangle \qquad V = \{U_{X}, U_{Y}, U_{Z} \}$$

$$V = \{X, Y, Z \}$$

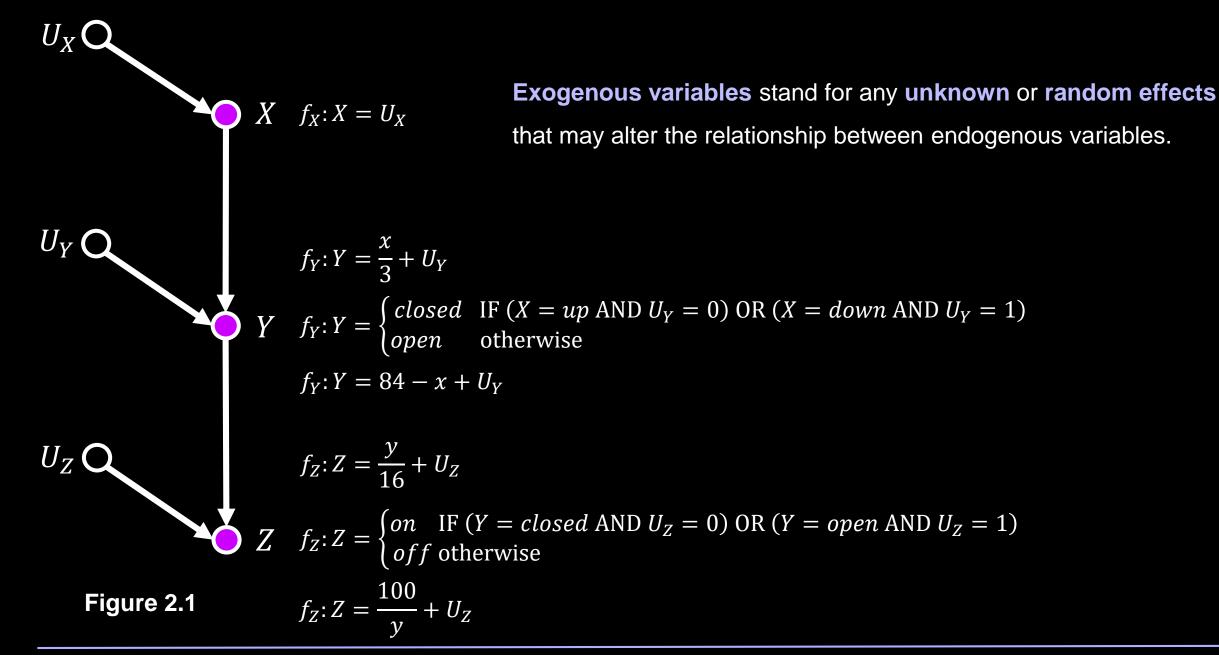
$$f_{Y} : Y = \{closed \ IF (X = up \ AND \ U_{Y} = 0) \ OR (X = down \ AND \ U_{Y} = 1)$$

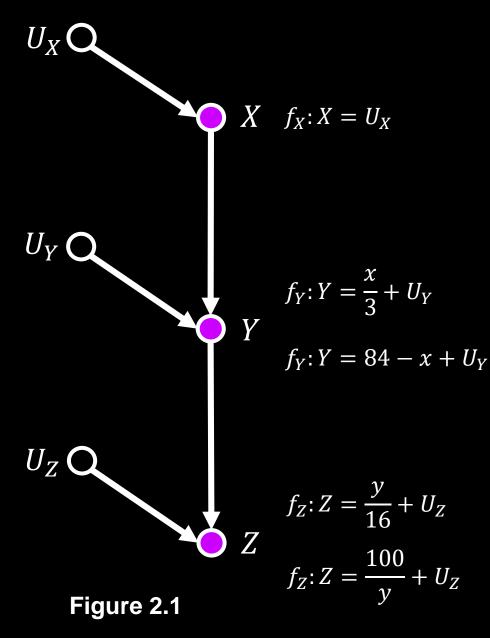
$$f_{Y} : Y = 84 - x + U_{Y}$$

$$f_{Z} : Z = \frac{V}{16} + U_{Z}$$

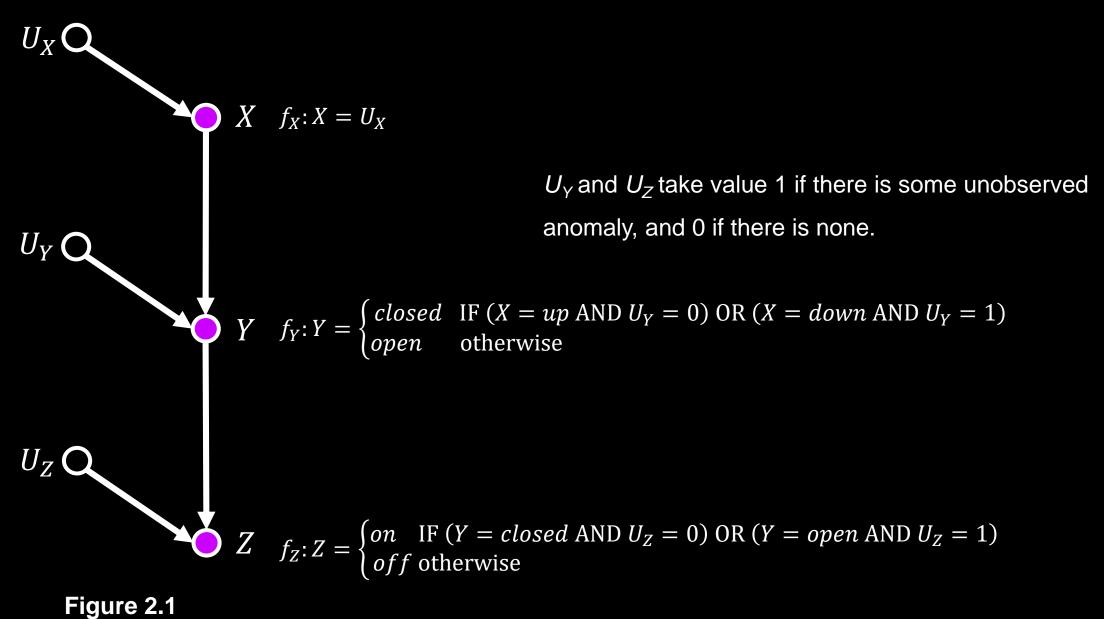
$$Z \quad f_{Z} : Z = \begin{cases} on \ IF (Y = closed \ AND \ U_{Z} = 0) \ OR (Y = open \ AND \ U_{Z} = 1) \end{cases}$$

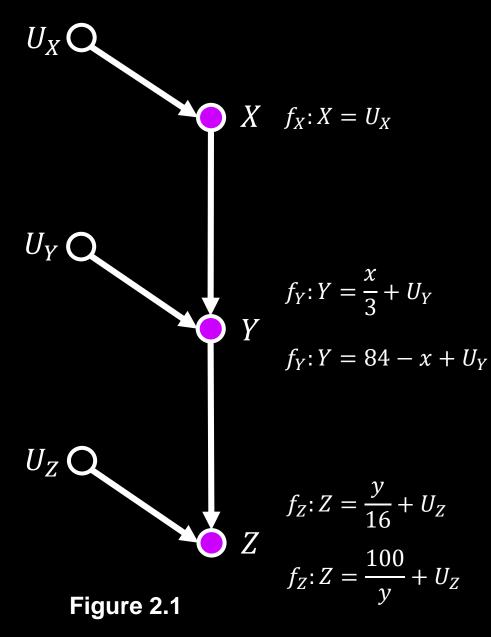
$$Figure 2.1 \qquad f_{Z} : Z = \frac{100}{y} + U_{Z}$$





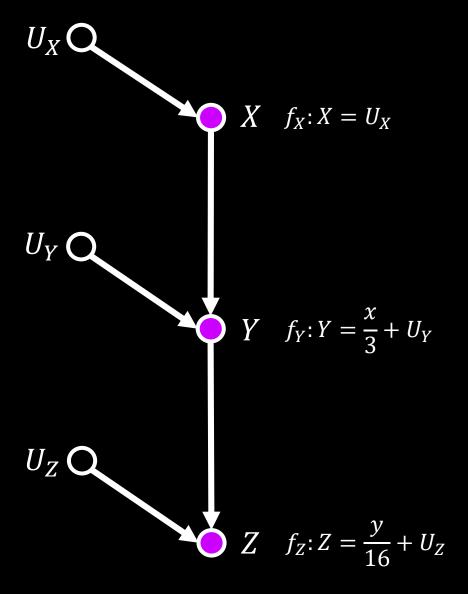
 U_{Y} and U_{Z} are additive factors that account for variations among individuals.





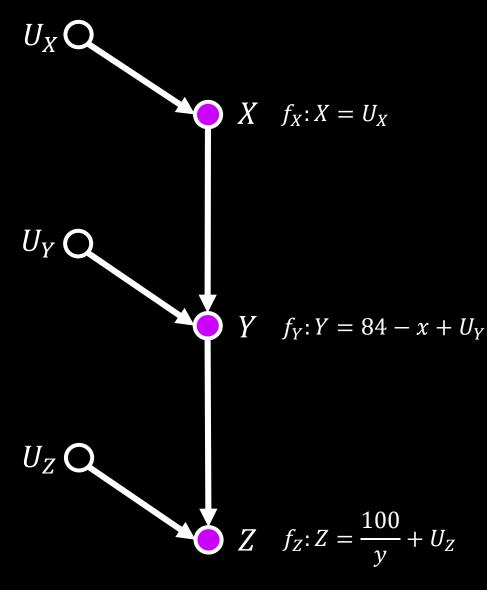
 M_1 and M_3 deal with continuous variables.

 M_2 deals with categorical variables.



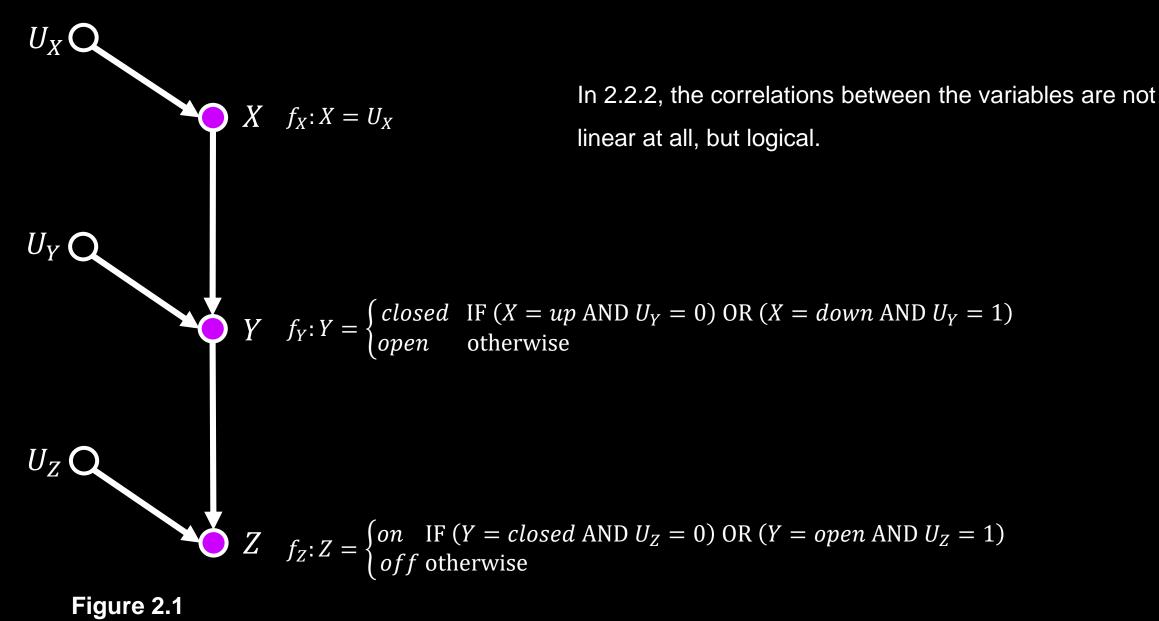
In 2.2.1, the relationships between variables are all positive, i.e.,

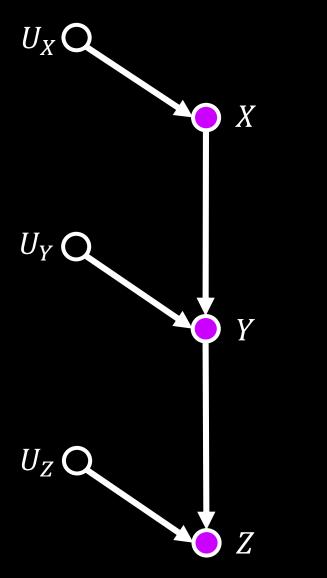
 the higher the value of the parent variable, the higher the values of the child variable.



In 2.2.3, for variables *Y* and *Z*, the correlation between them and their parents are all negative, i.e.,

 the higher the value of the parent variable, the lower the value of the child variable.





The three SCMs share no functions, except for f_X , but they share the same graphical structure.

The data sets generated by all three SCMs must share certain independencies, and we can predict those independencies simply by <u>examining the graphical model to the left.</u>

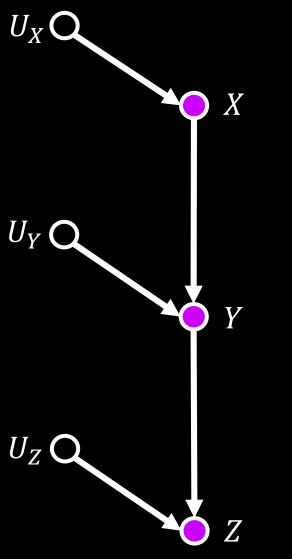


Figure 2.1

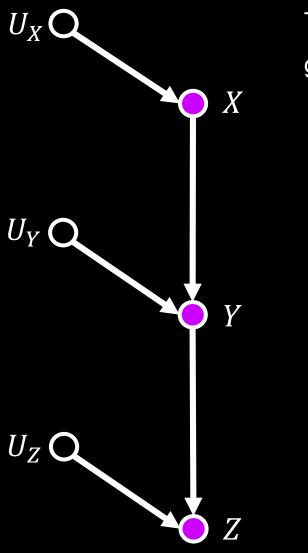
The independencies shared by data sets generated by these three SCMs, and the dependencies that are likely shared by all such SCMs are the following:

- **1.** Z and Y are likely dependent, i.e., for some pair of values *z*, *y*
 - $P(Z = z | Y = y) \neq P(Z = z)$
- 2. Y and X are likely dependent, i.e., for some pair of values y, x $P(Y = y | X = x) \neq P(Y = y)$
- **3.** Z and X are likely dependent, i.e., for some pair of values *z*, *x*

$$P(Z = z | X = x) \neq P(Z = z)$$

4. Z and X are independent, conditional on Y, i.e., for all values x, y, z

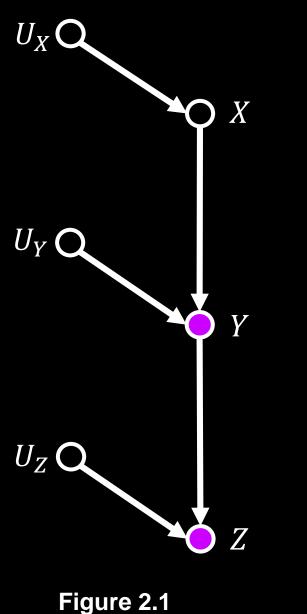
$$P(Z = z | X = x, Y = y) = P(Z = z | Y = y)$$



To understand why these independencies and dependencies hold, let's examine the graphical model.

Any two variables with an edge between them are likely dependent.

An arrow from one variable to another indicates that the first variable causes the second, that is, the value of the first variable is part of the function that determines the value of the second variable.



To understand why these independencies and dependencies hold, let's examine the graphical model.

Any two variables with an edge between them are likely dependent.

An arrow from one variable to another indicates that the first variable causes the second, that is, the value of the first variable is part of the function that determines the value of the second variable.

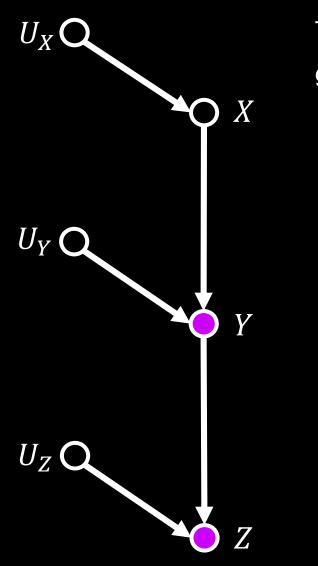
Y causes Z

Z depends on *Y*, there is some case in which changing the value of *Y* changes the value of *Z*.

$$I_Z: Z = \frac{y}{16} + U_Z$$

 $f_Z: Z = \begin{cases} on & \text{IF } (Y = closed \text{ AND } U_Z = 0) \text{ OR } (Y = open \text{ AND } U_Z = 1) \\ off \text{ otherwise} \end{cases}$

$$f_Z: Z = \frac{100}{y} + U_Z$$



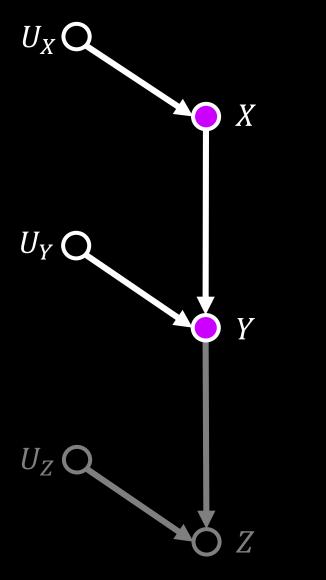
To understand why these independencies and dependencies hold, let's examine the graphical model.

• Any two variables with an edge between them are likely dependent.

When we examine those variables in the data set, the probability that one variable takes a given value will change, given that we know the value of the other variable.

The probability that Z takes value z, when we know that the value of the variable Y is equal to y, is different from the probability that Z takes the value z, when we do not know which value Y takes, i.e.,

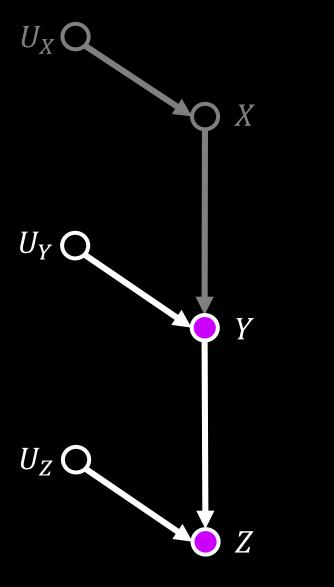
 $P(Z = z | Y = y) \neq P(Z = z)$



Therefore, we can conclude that in a typical causal model, regardless of the specific functions, two variables connected by an edge are likely dependent.

• X and Y are likely dependent

$$f_Y:Y=\frac{x}{3}+U_Y$$



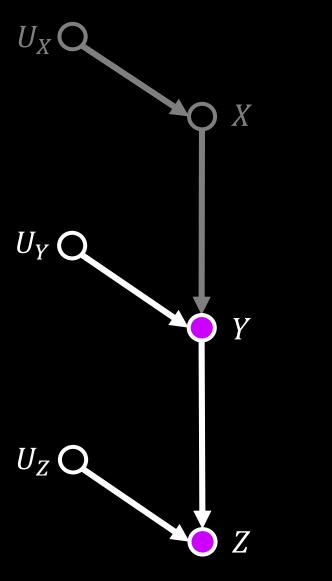
Therefore, we can conclude that in a typical causal model, regardless of the specific functions, two variables connected by an edge are likely dependent.

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$$f_Y: Y = \frac{x}{3} + U_Y$$

• Y and Z are likely dependent

$$f_Z: Z = \frac{y}{16} + U_Z$$



Therefore, we can conclude that in a typical causal model, regardless of the specific functions, two variables connected by an edge are likely dependent.

• X and Y are likely dependent

$$Y_Y:Y=\frac{x}{3}+U_Y$$

• Y and Z are likely dependent

$$f_Z: Z = \frac{y}{16} + U_Z$$

Why in general we say **likely dependent** and not simply **dependent**?

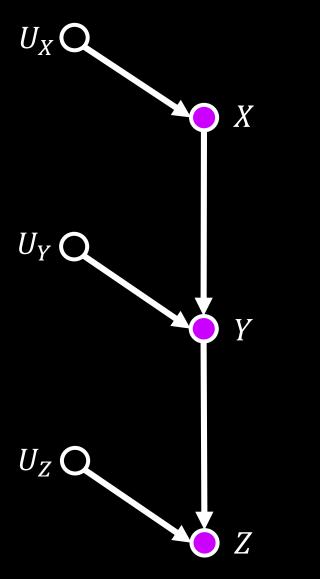


Figure 2.1

Therefore, we can conclude that in a typical causal model, regardless of the specific functions, two variables connected by an edge are likely dependent.

• X and Y are likely dependent

$$Y:Y=\frac{x}{3}+U_Y$$

• Y and Z are likely dependent

$$Z: Z = \frac{y}{16} + U_Z$$

Why in general we say **likely dependent** and not simply **dependent**?

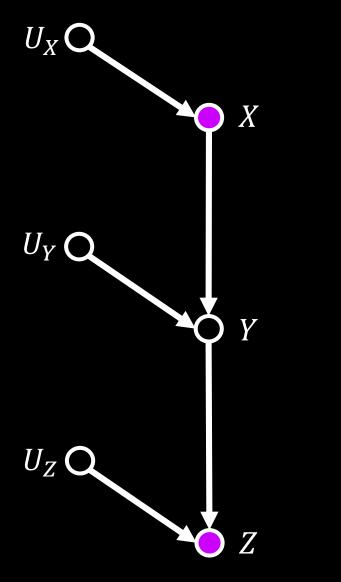
Consider X and U_Y be fair coins, and let Y = 1 if and only if $X = U_Y$, then

$$P(Y = 1 | X = 1) = P(Y = 1 | X = 0) = P(Y = 1) = \frac{1}{2}.$$

Such pathological cases require precise numerical probabilities to achieve independence

$$P(X = 1) = P(U_X = 1) = \frac{1}{2}$$

they are rare, and can be ignored for all practical purposes.



Furthermore, in a typical causal model, regardless of the specific functions, two variables connected by a directed path are likely dependent.

 $f_Y: Y = \frac{x}{3} + U_Y$ \downarrow $f_Z: Z = \frac{y}{16} + U_Z = \frac{\frac{x}{3} + U_Y}{16} + U_Z$

There are pathological cases in which the above is not true!!!

This the reason why we say likely dependent and not just dependent.

Figure 2.1

• X and Z are likely dependent

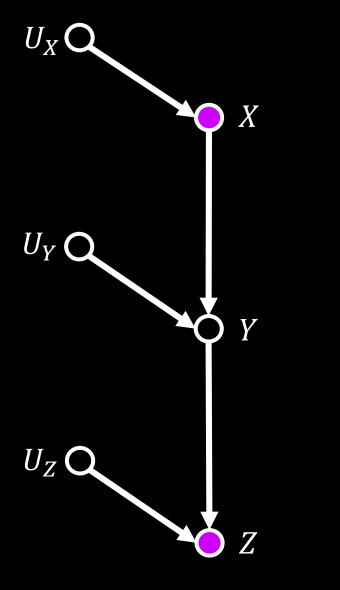


Figure 2.1

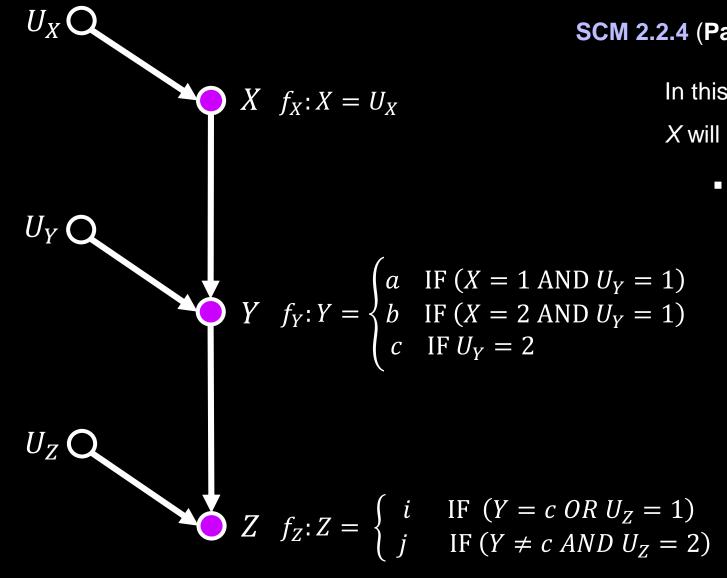
SCM 2.2.4 (Pathological Case of Intransitive Dependence)

$$M_4 = \langle U, V, F_4 \rangle$$

$$U = \{U_X, U_Y, U_Z\}$$
$$V = \{X, Y, Z\}$$
$$F_4 = \{f_X, f_Y, f_Z\}$$

$$U_X = \{1,2\} \qquad X = \{1,2\}$$
$$U_Y = \{1,2\} \qquad Y = \{a,b,c\}$$
$$U_Z = \{1,2\} \qquad Z = \{i,j\}$$

 $U_X \mathbf{Q}$ SCM 2.2.4 (Pathological Case of Intransitive Dependence) $X f_X: X = U_X$ $M_4 = \langle U, V, F_4 \rangle$ $U = \{U_X, U_Y, U_Z\}$ $U_Y \mathbf{O}$ $V = \{X, Y, Z\}$ Y $f_Y: Y = \begin{cases} a & \text{IF } (X = 1 \text{ AND } U_Y = 1) \\ b & \text{IF } (X = 2 \text{ AND } U_Y = 1) \\ c & \text{IF } U_Y = 2 \end{cases}$ $F_4 = \{f_X, f_Y, f_Z\}$ $U_X = \{1,2\}$ $X = \{1,2\}$ $U_Z \mathbf{Q}$ $U_Y = \{1, 2\}$ $Y = \{a, b, c\}$ $\sum_{i} \overline{f_{Z}: Z} = \begin{cases} i & \text{IF } (Y = c \ OR \ U_{Z} = 1) \\ j & \text{IF } (Y \neq c \ AND \ U_{Z} = 2) \end{cases}$ $U_Z = \{1, 2\}$ $Z = \{i, j\}$



SCM 2.2.4 (Pathological Case of Intransitive Dependence)

In this case, no matter what value U_Y and U_Z take, X will have no effect on the value that Z takes;

 changes in X account for variation in Y between a and b, but Y does not affect Z unless it takes value c.

Therefore, X and Z vary independently in this model.

Intransitive Case

By now it should be clear why previously we talked about **likely dependent** and not about dependent.

Figure 2.1

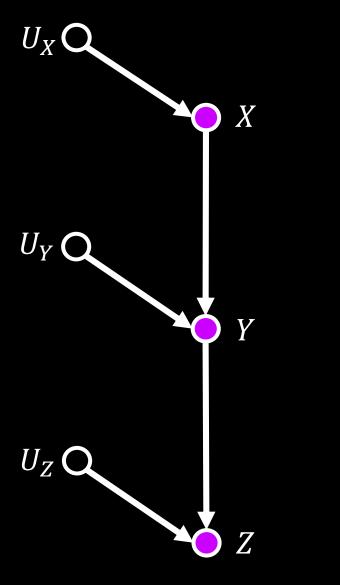


Figure 2.1

To summarize we have

• X and Y are likely dependent

$$f_Y:Y=\frac{x}{3}+U_Y$$

• Y and Z are likely dependent

$$f_Z: Z = \frac{y}{16} + U_Z$$

• X and Z are likely dependent

 $f_Z: Z = \frac{y}{16} + U_Z = \frac{\frac{x}{3} + U_Y}{16} + U_Z$

 $U_X \mathbf{Q}$ X $U_Y \mathbf{Q}$ $U_Z \mathbf{Q}$

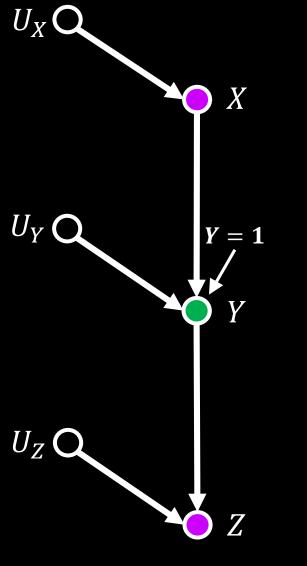
Now, let's consider the following point

• Z and X are independent, conditionally on Y, i.e., for all values x, y, z

$$P(Z = z | X = x, Y = y) = P(Z = z | Y = y)$$
 $Z \perp X | Y$

Remember that when we condition on *Y*, we filter the data into groups based on the value of *Y*.

		U _x	U,	U _z	X	Y	Ζ	-	U _x	U,	U _z	X	Y	Ζ
	V	1	0	0	1	1	1	<u>-</u>	1	0	0	1	1	1
\mathbf{Y}	Y	-	0	0	1	1	1	cutoring	1	0	0	1	1	1 2
		1	0	1	1	1	2	Table filtering	2	1	0	2	1	1
		2	0	0	2	2	2	$\frac{1215}{55}$	2	1	0	2	1	1
		1	-1	0	1	2	2		2	1	1	2	1	2
\cap		1	-1	0	1	2	2	Table fill						
		2	1	0	2	1	1	Table filtering	U _x	U,	U _z	X	Ŷ	Ζ
		2	1	0	2	1	1	by $\gamma = 2$	2	0	0	2	2	2
	7	2	1	1	2	1	2		1	-1	0	1	2	2
	L	2	0	0		1 2			1	-1	0	1	2	2
		Z	0	0	2	2	2	-	2	0	0	2	2	2
Figure 2.1	X = U	$J_{\mathbf{v}}$:	Y =	X -	$U_{\mathbf{v}}$:	Z	= Y -	+ <i>U_Z</i> ;						



We compare *X* and *Z* on all the cases where Y = 1, and on all the cases where Y = 2. Let's assume that we are looking at the cases where Y = 1. We want to know whether, in these cases only, the value of *Z* is independent of the value of *X*.

Previously, we determined that X and Z are likely dependent, because when the value of X changes, the value of Y likely changes, and when the value of Y changes, the value of Z is likely to change.

U _x	U _y	U _z	X	Y	Ζ
1	0	0	1	1	1
1	0	0	1	1	1
1	0	1	1	1	2
2	1	0	2	1	1
2	1	0	2	1	1
2	1	1	2	1	2

Figure 2.1

X

Y = 1

 $U_X \mathbf{Q}$

 $U_Y \mathbf{O}$

 $U_Z \mathbf{O}$

We compare *X* and *Z* on all the cases where Y = 1, and on all the cases where Y = 2. Let's assume that we are looking at the cases where Y = 1. We want to know whether, in these cases only, the value of *Z* is independent of the value of *X*.

$$P(Z = z | X = x, Y = y) = P(Z = z | Y = y), \quad \forall x, z \; Y = 1$$

U _x	U _y	U _z	X	Y	Ζ
1	0	0	1	1	1
1	0	0	1	1	1
1	0	1	1	1	2
2	1	0	2	1	1
2	1	0	2	1	1
2	1	1	2	1	2

Figure 2.1

Causal Networks: Learning and Inference

Fabio Stella and Luca Bernardinello

2019 September 23rd - 27th

 $U_X \mathbf{Q}$

 $U_Y \mathbf{Q}$

 $U_Z \mathbf{Q}$

We compare X and Z on all the cases where Y = 1, and on all the cases where Y = 2. Let's assume that we are looking at the cases where Y = 1. X We want to know whether, in these cases only, the value of Z is independent of the value of X. $P(Z = z | X = x, Y = y) = P(Z = z | Y = y), \quad \forall x, z \; Y = 1$ Y = 1 $P(Z = 1 | X = 1, Y = 1) = \frac{2}{3}$ $P(Z = 2|X = 1, Y = 1) = \frac{1}{3}$ $\boldsymbol{U}_{\mathbf{x}}$ Uv U, Ζ X Y 1 0 0 1 1 1 1 1 2 Ζ 2 1 1 <u>2</u> 1 0 **2** 1 1 Figure 2.1 2 2 1 2

 $U_X \mathbf{Q}$

 $U_Y \mathbf{Q}$

 $U_Z \mathbf{O}$

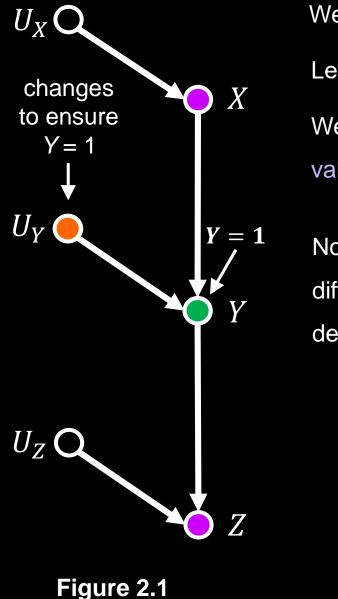
We compare X and Z on all the cases where Y = 1, and on all the cases where Y = 2. Let's assume that we are looking at the cases where Y = 1. X We want to know whether, in these cases only, the value of Z is independent of the value of X. $P(Z = z | X = x, Y = y) = P(Z = z | Y = y), \quad \forall x, z \; Y = 1$ Y = 1*V* $P(Z = 1 | X = 1, Y = 1) = \frac{2}{3}$ $P(Z = 1 | X = 2, Y = 1) = \frac{2}{3}$ $P(Z = 2|X = 1, Y = 1) = \frac{1}{3}$ $P(Z = 2|X = 2, Y = 1) = \frac{1}{3}$ $U_x \quad U_y \quad U_z \quad X \quad Y$ Ζ 1 0 0 1 1 Figure 2.1 2 1 2

 $U_X \mathbf{Q}$

 $U_Y \mathbf{Q}$

 $U_Z \mathbf{Q}$

We compare X and Z on all the cases where Y = 1, and on all the cases where Y = 2. Let's assume that we are looking at the cases where Y = 1. X We want to know whether, in these cases only, the value of Z is independent of the value of X. $P(Z = z | X = x, Y = y) = P(Z = z | Y = y), \quad \forall x, z \; Y = 1$ Y = 1 $P(Z = 1 | X = 1, Y = 1) = \frac{2}{3} \quad P(Z = 1 | X = 2, Y = 1) = \frac{2}{3} \quad P(Z = 1 | Y = 1) = \frac{4}{6} = \frac{2}{3}$ $P(Z = 2|X = 1, Y = 1) = \frac{1}{3}$ $P(Z = 2|X = 2, Y = 1) = \frac{1}{3}$ $P(Z = 2|Y = 1) = \frac{2}{6} = \frac{1}{3}$ U_{y} U_{z} XU_x Ζ 1 1 0 0 1 1 0 0 1 1 1 0 1 **1 1** 2 1 0 **2 1** 2 Y = 1<u>2 1 0 **2 1**</u> Figure 2.1 2 1 2 2



We compare *X* and *Z* on all the cases where Y = 1, and on all the cases where Y = 2. Let's assume that we are looking at the cases where Y = 1. We want to know whether, in these cases only, the value of *Z* is independent of the value of *X*.

Now, however, examining only the cases where Y = 1, when we select cases with different values of *X*, the value of U_Y changes so as to keep *Y* at Y = 1, but since *Z* depends only on *Y* and U_Z , not on U_Y , the value of *Z* remains unaltered.

U _x	U _y	U _z	X	Y	Ζ
1	0	0	1	1	1
1	0	0	1	1	1
1	0	1	1	1	2
2	1	0	2	1	1
2	1	0	2	1	1
2	1	1	2	1	2

 $Y = X - U_V$

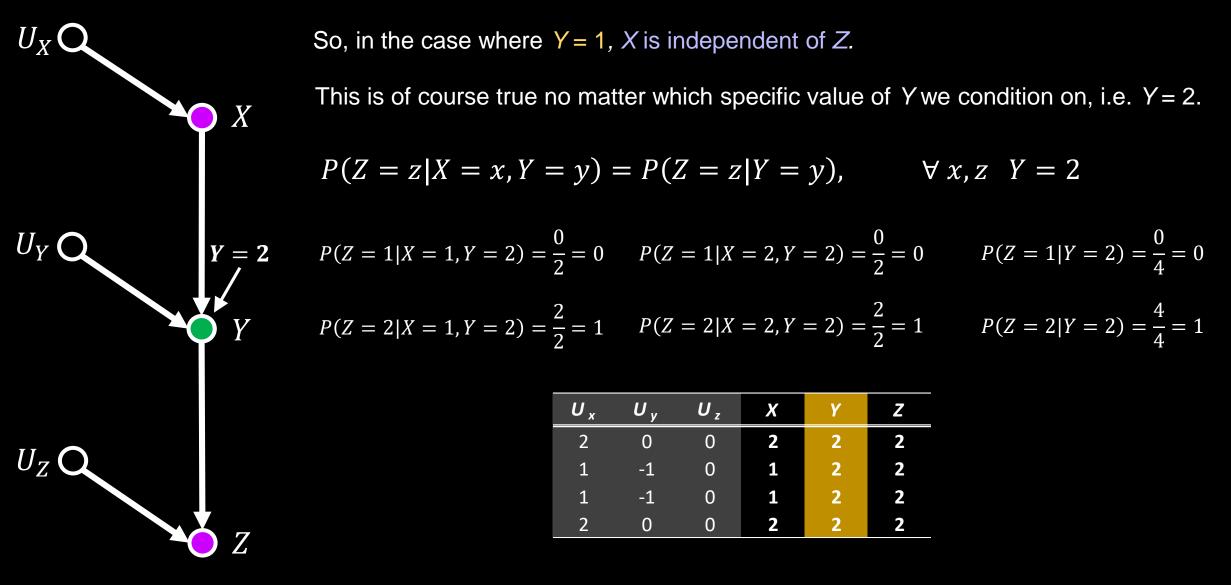


Figure 2.1

Therefore, X is independent of Z, conditionally on Y.

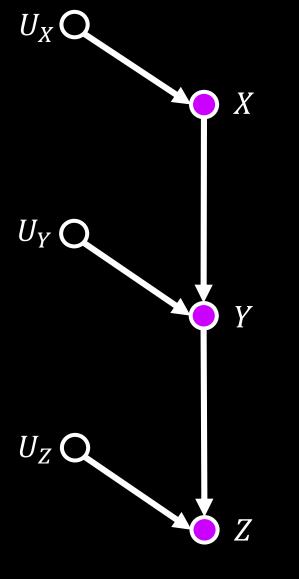


Figure 2.1

This configuration of variables, three nodes and two edges, with one edge directed into and one edge directed out of the middle variable, is called a **chain**.

In any graphical model, given any two variables *X* and *Y*, if the only path between *X* and *Y* is composed entirely of chains, then *X* and *Y* are Independent conditional on any intermediate variable on the path.

This independence relation holds regardless of the functions that connect the variables.

Rule 1 (Conditional Independence in Chains)

Two variables, X and Y, are conditionally independent given Z, if there is only one unidirectional path between X and Y, and Z is any set of variables that intercepts that path.

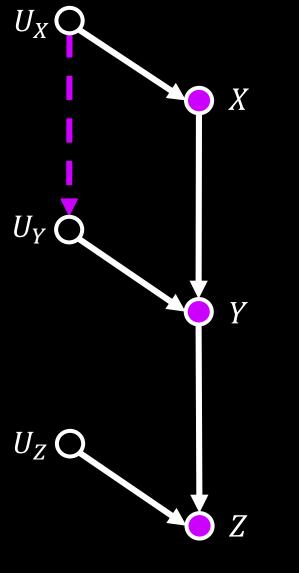


Figure 2.1

If, for instance, U_X were a cause of U_Y (dashed arrow in **Figure 2.1**), then conditioning on *Y* would not necessarily make *X* and *Z* independent, because variations in *X* could still be associated with variations in *Y*, through their error terms.

Rule 1 only holds when we assume that the error terms U_X , U_Y , and U_Z are independent of each other.

The previous, gives the following rule: WARNING \longrightarrow N, Y and Z in Rule 1 do not refer to Figure 2.1.

Rule 1 (Conditional Independence in Chains)

Two variables, X and Y, are conditionally independent given Z, if there is only one unidirectional path between X and Y, and Z is any set of variables that intercepts that path.

Now, consider the graphical model in Figure 2.2.

This structure might represent, for example, the causal mechanism that connects a day's temperature in a city in degrees Fahrenheit (X), the number of sales at a local ice cream shop on that day (Y), and the number of violent crimes in the city on that day (Z).

SCM 2.2.5 (Temperature, Ice Cream Sales, and Crime) U_X $M = \langle U, V, F \rangle$ $f_X: X = U_X$ X $U = \{U_X, U_Y, U_Z\}$ U_{Y} C U_Z $V = \{X, Y, Z\}$ $f_Z: Z = \frac{x}{10} + U_Z$ $f_Y: Y = 4x + U_Y$ $F_3 = \{f_X, f_Y, f_Z\}$ Figure 2.2

Now, consider the graphical model in Figure 2.2.

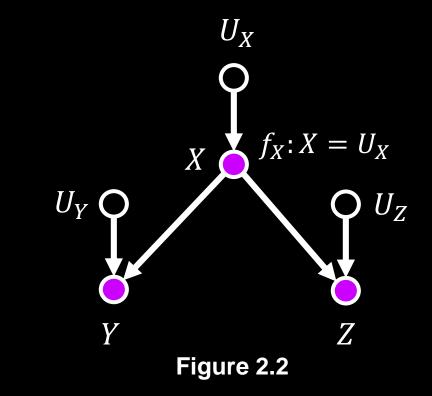
Causal mechanism that connects the state (*up* or *down*) of a switch (*X*), the state (*on* or *off*) of one light bulb (*Y*), and the state (*on* or *off*) of a second light bulb (*Z*).

SCM 2.2.6 (Switch and Two Light Bulbs)

 $M = \langle U, V, F \rangle$

 $f_Y: Y = \begin{cases} on & \text{IF } (X = up \text{ AND } U_Y = 0) \text{ OR } (X = down \text{ AND } U_Y = 1) \\ off & \text{otherwise} \end{cases}$

$$f_Z: Z = \begin{cases} on & \text{IF} (X = up \text{ AND } U_Z = 0) \text{ OR } (X = down \text{ AND } U_Z = 1) \\ off & \text{otherwise} \end{cases}$$



If we assume that the error terms U_X , U_Y , and U_Z are independent, then by examining the graphical model in **Figure 2.2**, we can determine that the **SCMs 2.2.5/2.2.6** share the following dependencies and independencies:

1. X and Y are likely dependent, i.e., for some pair of values x, y

$$P(X = x | Y = y) \neq P(X = x)$$

2. X and Z are likely dependent, i.e., for some pair of values x, z

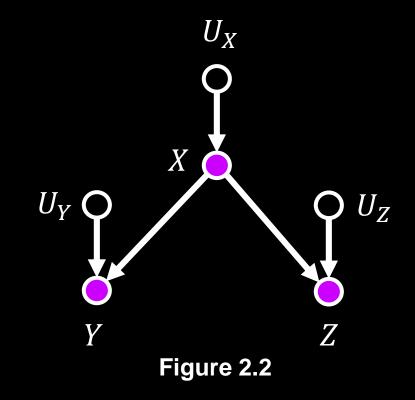
$$P(X = x | Z = z) \neq P(X = x)$$

3. Z and Y are likely dependent, i.e., for some pair of values z, y

$$P(Z = z | Y = y) \neq P(Z = z)$$

4. Y and Z are independent, conditional on X, i.e., for all values x, y, z

$$P(Y = y | X = x, Z = z) = P(Y = y | X = x)$$



If we assume that the error terms U_X , U_Y , and U_Z are independent, then by examining the graphical model in **Figure 2.2**, we can determine that the **SCMs 2.2.5/2.2.6** share the following dependencies and independencies:

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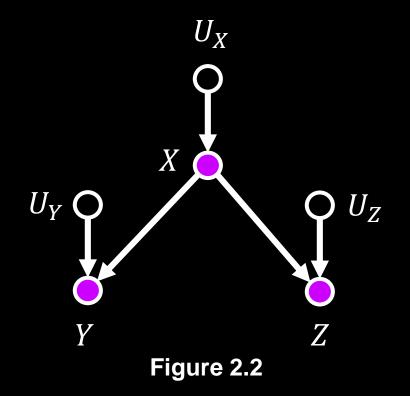
 $P(X = x | Y = y) \neq P(X = x)$

2. X and Z are likely dependent, i.e., for some pair of values x, z

 $P(X = x | Z = z) \neq P(X = x)$

Follow, once again, by the fact that Y and Z are both directly connected to X by an arrow.

When the value of X changes, the values of both Y and Z likely change.



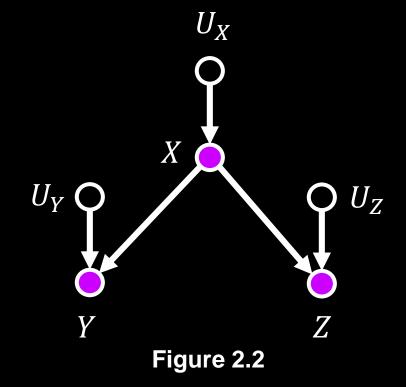
If we assume that the error terms U_X , U_Y , and U_Z are independent, then by examining the graphical model in **Figure 2.2**, we can determine that the **SCMs 2.2.5/2.2.6** share the following dependencies and independencies:

This tells us something further, however: If Y changes when X changes, and Z changes when X changes, then it is likely (though not certain) that Y changes together with Z, and vice versa.

Therefore, since a change in the value of Y gives us information about an associated change in the value of Z, Y and Z are likely dependent variables.

3. Z and Y are likely dependent, i.e., for some pair of values z, y

$$P(Z = z | Y = y) \neq P(Z = z)$$



If we assume that the error terms U_X , U_Y , and U_Z are independent, then by examining the graphical model in **Figure 2.2**, we can determine that the **SCMs 2.2.5/2.2.6** share the following dependencies and independencies:

Why, then, are Y and Z independent conditional on X? What happens when we condition on X?

We filter the data based on the value of X. So now, we are only comparing

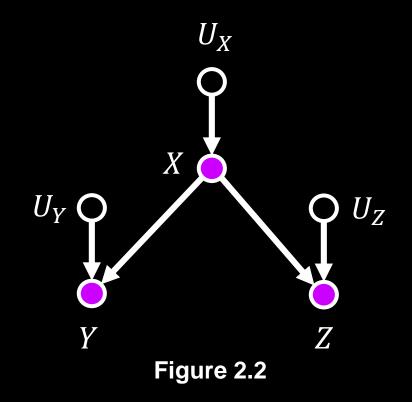
cases where the value of X is constant.

Since *X* does not change, the values of *Y* and *Z* do not change in accordance with it, they change only in response to U_Y and U_Z , which we have assumed independent.

Therefore, any additional changes in the values of Y and Z must be independent of each other.

4. Y and Z are independent, conditional on X, i.e., for all values x, y, z

P(Y = y | X = x, Z = z) = P(Y = y | X = x)



This configuration of variables, three nodes, with two arrows emanating from the middle variable, is called a **fork**.

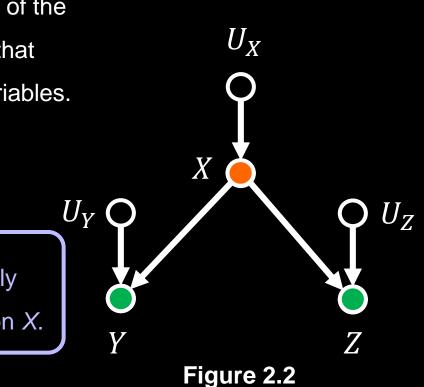
The middle variable in a fork (X) is a common cause of the other two variables (Y and Z), and of any of their descendants.

If two variables share a common cause, and if that common cause is part of the only path between them, then analogous reasoning to the above tells us that these dependencies and conditional independencies are true of those variables.

Therefore, we come by another rule:

Rule 2 (Conditional Independence in Forks)

If a variable *X* is a common cause of variables *Y* and *Z*, and there is only one path between *Y* and *Z*, then *Y* and *Z* are independent conditional on *X*.



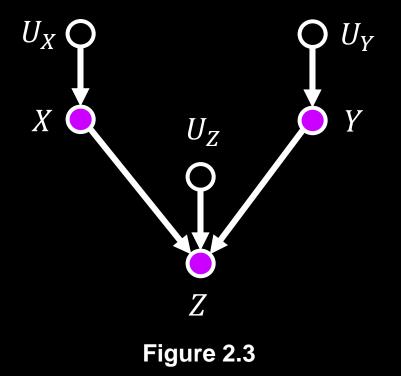
So far we have looked at two simple configurations of edges and nodes that can occur on a path between two variables: chains and forks.

There is a third such configuration that we speak of separately, because it carries with it unique considerations and challenges.

The third configuration contains a **collider** node, and it occurs when one node receives edges from two other nodes.

The simplest graphical causal model containing a **collider** is illustrated in **Figure 2.3**, representing a common effect *Z*, of two causes *X* and *Y*.

As is the case with every graphical causal model, all **SCMs** that have **Figure 2.3** as their graph share a set of dependencies and independencies that we can determine from the graphical model alone.



Assume that the error terms U_X , U_Y , and U_Z are independent, then by examining the graphical model in **Figure 2.3**, we can determine that any corresponding **SCMs** share the following dependencies and independencies:

1. X and Z are likely dependent, i.e., for some pair of values x, z

 $P(X = x | Z = z) \neq P(X = x)$

2. Y and Z are likely dependent, i.e., for some pair of values y, z

$$P(Y = y | Z = z) \neq P(Y = y)$$

3. X and Y are independent, i.e., for all pairs of values x, y

P(X = x | Y = y) = P(X = x)

4. X and Y are likely dependent, conditional on Z, i.e., for some values x, y, z

$$P(X = x | Y = y, Z = z) \neq P(X = x | Z = z)$$

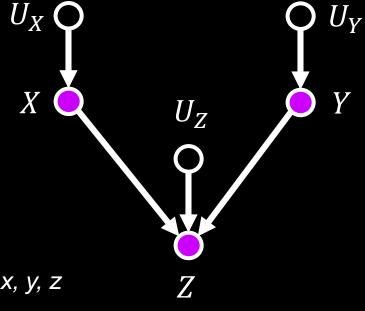


Figure 2.3

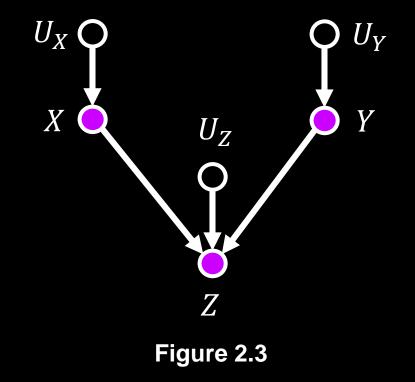
The truth of the first two points was established in Section 2.2 Chain and Forks.

1. X and Z are likely dependent, i.e., for some pair of values x, z

 $P(X = x | Z = z) \neq P(X = x)$

2. Y and Z are likely dependent, i.e., for some pair of values y, z

 $P(Y = y | Z = z) \neq P(Y = y)$



The third point is self-evident,

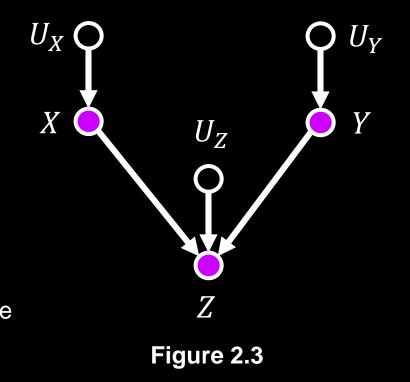
- neither X nor Y is a descendant or an ancestor of the other,
- nor do they depend for their value on the same variable,
- X and Y respond only to U_X and U_Y, which we assumed to be independent,

so there is no causal mechanism by which variations in the value of X should be associated with variations in the value of Y.

3. *X* and *Y* are independent, i.e., for all pairs of values *x*, *y*

P(X = x | Y = y) = P(X = x)

This independence also reflects our understanding of how causation operates in time; events that are independent in the present do not become dependent merely because they may have common effects in the future.



Why, then, does Point 4 hold?

Why would two independent variables suddenly become dependent when we condition on their common effect?

To answer this question, we return again to the definition of conditioning as filtering by the value of the conditioning variable.

When we condition on Z, we limit our comparison to cases in which Z takes the same value.

But remember that Z depends, for its value, on X and Y. So, when comparing cases where Z takes some value, any change in value of X must be compensated for by a change in the value of Y, otherwise, the value of Zwould change as well.

4. X and Y are likely dependent, conditional on Z, i.e., for some values x, y, z

$$P(X = x | Y = y, Z = z) \neq P(X = x | Z = z)$$

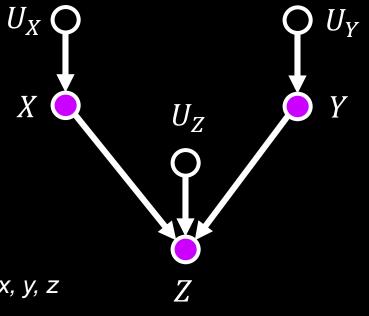
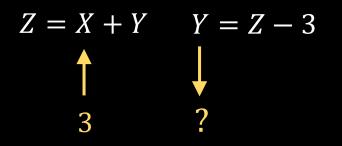


Figure 2.3

Why, then, does Point 4 hold?

Why would two independent variables suddenly become dependent when we condition on their common effect?

The reasoning behind this attribute of colliders, that conditioning on a collider node produces a dependence between the node's parents, can be difficult to grasp at first.



From the X value we learn nothing about the Y value, because the two numbers are independent

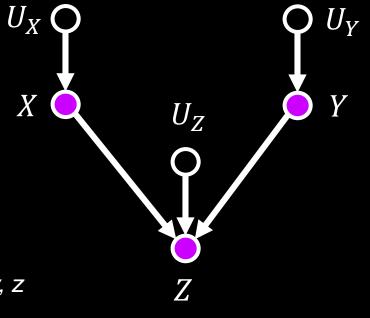


Figure 2.3

4. X and Y are likely dependent, conditional on Z, i.e., for all values x, y, z

$$P(X = x | Y = y, Z = z) \neq P(X = x | Z = z)$$

2019 September 23rd - 27th

Fabio Stella and Luca Bernardinello

Why, then, does Point 4 hold?

Why would two independent variables suddenly become dependent when we condition on their common effect?

The reasoning behind this attribute of colliders, that conditioning on a collider node produces a dependence between the node's parents, can be difficult to grasp at first.

$$Z = X + Y \qquad Y = 10 - 3$$

$$\uparrow \qquad \uparrow \qquad \downarrow$$

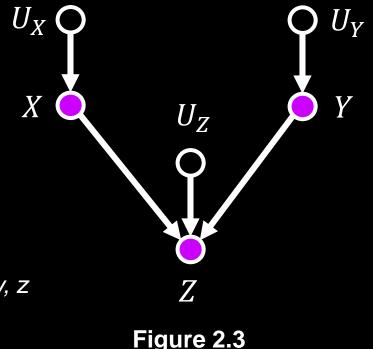
$$10 \quad 3 \qquad 7$$

Thus X and Y are dependent, given that Z = 10.

We are implicitly assuming U_{χ} , U_{γ} , and U_{γ} to be zero.

4. X and Y are likely dependent, conditional on Z, i.e., for all values x, y, z

$$P(X = x | Y = y, Z = z) \neq P(X = x | Z = z)$$





Consider a simultaneous (independent) toss of two fair

coins and a bell that rings whenever at least one of the

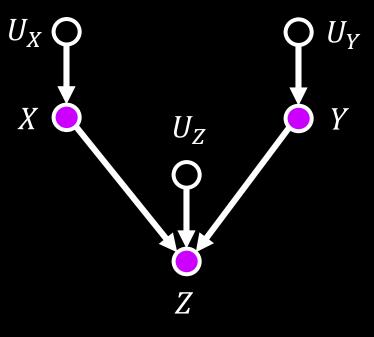
coin lands on heads.

 $X, Y \in \{heads, tails\}$ $Z \in \{silence, rings\}$

We know the following:

coin 1 landed on heads (X = heads)

Tells us nothing about the outcome of the toss of coin 2 (*Y*)





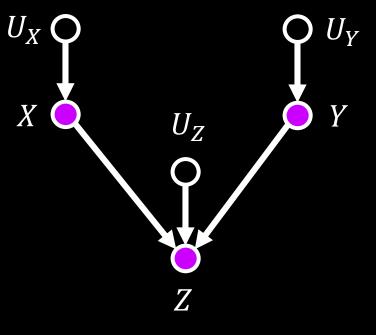
 $X, Y \in \{heads, tails\}$ $Z \in \{silence, rings\}$

We know the following:

- we hear the **bell ringing** (Z = rings)
- coin 1 landed on tails (X = tails)

Tells us that coin 2 must have landed on heads, i.e., Y = heads.

coin lands on heads.



Consider a simultaneous (independent) toss of two fair

coins and a bell that rings whenever at least one of the





We know the following:

- we hear the **bell ringing** (Z = rings)
- we know that coin 2 landed on heads (Y = heads)

 $P(X = heads | Z = rings) \neq P(X = heads | Y = heads, Z = rings)$

Consider a simultaneous (independent) toss of two fair

coins and a bell that rings whenever at least one of the

coin lands on heads.

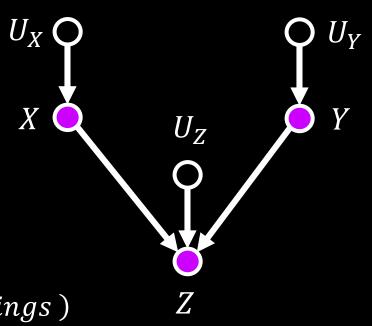


Figure 2.3

To see the latter calculation, consider the initial probabilities in **Table 2.1**. We see that

P(X = heads) = 0.5

P(X = tails) = 0.5

Table 2.1 Probability distribution for two flips of a fair coin, with Xrepresenting flip one, Y representing flip two, and Z representing a bellthat rings if either flip results in heads.

X	Y	Ζ	
Coin 1	Coin 2	Bell	P(X, Y, Z)
heads	heads	rings	0.25
heads	tails	rings	0.25
tails	heads	rings	0.25
tails	tails	silence	0.25

To see the latter calculation, consider the initial probabilities in **Table 2.1**. We see that

P(X = heads) = 0.5 = P(X = heads|Y = tails)

P(X = tails) = 0.5

Table 2.1 Probability distribution for two flips of a fair coin, with X representing flip one, Y representing flip two, and Z representing a bell that rings if either flip results in heads.

X	Y	Ζ	P(X, Y, Z)	
Coin 1	Coin 2	Bell	Ρ(Λ, Ι, Ζ)	
heads	tails	rings	0.25	
tails	tails	silence	0.25	

To see the latter calculation, consider the initial probabilities in **Table 2.1**. We see that

P(X = heads) = 0.5 = P(X = heads|Y = tails) = P(X = heads|Y = heads)

P(X = tails) = 0.5

Table 2.1 Probability distribution for two flips of a fair coin, with X representing flip one, Y representing flip two, and Z representing a bell that rings if either flip results in heads.

X Coin 1	Y Coin 2	Z	P(X, Y, Z)
Coin 1 heads	Coin 2 heads	Bell rings	0.25
tails	heads	rings	0.25

To see the latter calculation, consider the initial probabilities in **Table 2.1**. We see that

P(X = heads) = 0.5 = P(X = heads|Y = tails) = P(X = heads|Y = heads)

P(X = tails) = 0.5 = P(X = tails|Y = tails)

Table 2.1 Probability distribution for two flips of a fair coin, with X representing flip one, Y representing flip two, and Z representing a bell that rings if either flip results in heads.

X	Y	Ζ	P(X, Y, Z)	
Coin 1	Coin 2	Bell	Ρ(Λ, Ι, Ζ)	
heads	tails	rings	0.25	
tails	tails	silence	0.25	

To see the latter calculation, consider the initial probabilities in **Table 2.1**. We see that

P(X = heads) = 0.5 = P(X = heads|Y = tails) = P(X = heads|Y = heads)

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X	Y	Ζ	P(X, Y, Z)
Coin 1	Coin 2	Bell	Γ(Λ, Ι, Ζ)
heads	heads	rings	0.25
tails	heads	rings	0.25

To see the latter calculation, consider the initial probabilities in **Table 2.1**. We see that

P(X = heads) = 0.5 = P(X = heads|Y = tails) = P(X = heads|Y = heads)

P(X = tails) = 0.5 = P(X = tails|Y = tails) = P(X = tails|Y = heads)

X and Y are independent

Table 2.1 Probability distribution for two flips of a fair coin, with X representing flip one, Y representing flip two, and Z representing a bell that rings if either flip results in heads.

X	Y	Ζ	
Coin 1	Coin 2	Bell	P(X, Y, Z)
heads	heads	rings	0.25
heads	tails	rings	0.25
tails	heads	rings	0.25
tails	tails	silence	0.25

Now let's condition on Z = rings and Z = silence, the resulting data subsets are shown in **Table 2.2**.

Table 2.2 Condi	tional probability distribution f	or the
distribution in Ta	ble 2.1.	
V	V	

X	Y	P(X, Y Z=silence)
Coin 1	Coin 2	F(X, T Z=Silence)
heads	heads	0
heads	tails	0
tails	heads	0
tails	tails	1
X	Y	D(V, V Z - vinge)
Coin 1	Coin 2	P(X, Y Z=rings)
heads	heads	0.333
heads	tails	0.333
tails	heads	0.333

P(X = heads | Z = rings) =

Now let's condition on Z = rings and Z = silence, the resulting data subsets are shown in **Table 2.2**.

$$P(X = heads | Z = rings) =$$

X Coin 1	Y Coin 2	P(X, Y Z=rings)
heads	heads	0.333
heads	tails	0.333
tails	heads	0.333
tails	tails	0

Now let's condition on Z = rings and Z = silence, the resulting data subsets are shown in **Table 2.2**.

$$P(X = heads | Z = rings) = \frac{1}{3} + \frac{1}{3} = \frac{2}{3}$$

P(X = heads | Y = heads, Z = rings) =

X Coin 1	γ Coin 2	P(X, Y Z=rings)
heads	heads	0.333
heads	tails	0.333
tails	heads	0.333
tails	tails	0

Now let's condition on Z = rings and Z = silence, the resulting data subsets are shown in **Table 2.2**.

$$P(X = heads | Z = rings) = \frac{1}{3} + \frac{1}{3} = \frac{2}{3}$$

$$P(X = heads | Y = heads, Z = rings) = \frac{1}{2}$$

X Coin 1	Y Coin 2	P(X, Y Z=rings)
heads	heads	0.333
heads	tails	0.333
tails	heads	0.333
tails	tails	0

Now let's condition on Z = rings and Z = silence, the resulting data subsets are shown in **Table 2.2**.

Given Z = rings, the probability of X = heads is $P(X = heads | Z = rings) = \frac{1}{3} + \frac{1}{3} = \frac{2}{3}$

However, when we learn that Y = heads, the probability of X = heads changes as follows

$$P(X = heads | Y = heads, Z = rings) = \frac{1}{2}$$

$$P(X = heads | Y = heads, Z = rings) \neq P(X = heads | Z = rings)$$

X Coin 1	Y Coin 2	P(X, Y Z=rings)
heads	heads	0.333
heads	tails	0.333
tails	heads	0.333
tails	tails	0

Therefore, we conclude that *X* and *Y* are dependent given Z = rings.

Table 2.2 Conditional probability distribution for thedistribution in Table 2.1.

X	Y	P(X, Y Z=silence)	
Coin 1	Coin 2	$P(\Lambda, T Z-SHETCE)$	
heads	heads	0	
heads	tails	0	
tails	heads	0	
tails	tails	1	

A more pronounced dependence occurs, of course, when the bell does not ring (Z = silence), because then we know that both coins must have landed on tails.

Just as conditioning on a collider makes previously independent variables

dependent, so too does conditioning on any descendant of a collider.

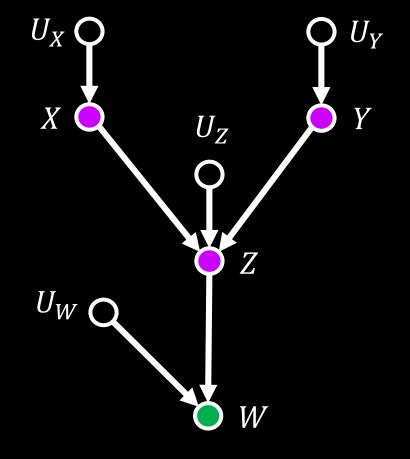


X Y Z

To see why this is true, let's return to our example of two independent coins and a bell.

Suppose we do not hear the bell directly, but instead rely on a witness (*W*) who is somewhat unreliable;

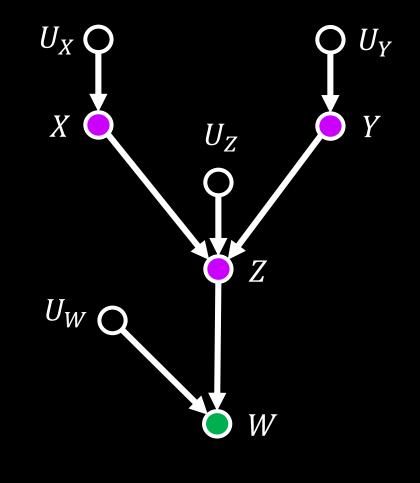
- whenever the bell does not ring (Z = silence), 50% chance that the witness will falsely report that it did (W = 1), and 50% chance that will correctly report that it did not (W = 0).
- whenever the bell rings (Z = rings), 100% chance that the witness will report that it did (W = 1).



Probabilities for all combinations of *X*, *Y* and *W* are shown in **Table 2.3**.

Table 2.3 Probability distribution for two flips of a fair coin and a bell that rings if either flip results in heads, with *X* representing flip one, *Y* representing flip two, and *W* representing a witness who, with variable reliability, reports whether or not the bell has rung.

X	Y	W	P(X, Y, W)
Coin 1	Coin 2	witness	Γ(Λ, Ι, νν)
heads	heads	1	0.250
heads	tails	1	0.250
tails	heads	1	0.250
tails	tails	1	0.125
tails	tails	0	0.125



How do we get Table 2.3 ?

Table 2.1 Probability distribution for two flips of a fair coin, with X representing flip one, Y representing flip two, and Z representing a bell that rings if either flip results in heads.

Table 2.3 Probability distribution for two flips of a fair coin and a bell that rings if either flip results in heads, with *X* representing flip one, *Y* representing flip two, and *W* representing a witness who, with variable reliability, reports whether or not the bell has rung.

X	Y	Ζ	P(X, Y, Z)	X	Y	W	P(X, Y, W)
Coin 1	Coin 2	Bell	Γ(Λ, Ι, Ζ)	Coin 1	Coin 2	witness	Γ(Λ, Ι, νν)
heads	heads	rings	0.25	heads	heads	1	0.250
heads	tails	rings	0.25	heads	tails	1	0.250
tails	heads	rings	0.25	tails	heads	1	0.250
tails	tails	silence	0.25	tails	tails	1	0.125
				tails	tails	0	0.125

How do we get Table 2.3 ?

Table 2.1 Probability distribution for two flips of a fair coin, with X representing flip one, Y representing flip two, and Z representing a bell that rings if either flip results in heads.

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X	Y	Ζ	P(X, Y, Z)	X	Y	W	P(X, Y, W)
Coin 1	Coin 2	Bell	Ρ(Λ, Ι, Ζ)	Coin 1	Coin 2	witness	$P(\Lambda, T, VV)$
heads	heads	rings	0.25	heads	heads	1	0.250
heads	tails	rings	0.25	heads	tails	1	0.250
tails	heads	rings	0.25	tails	heads	1	0.250
_				_		4	
		0					
						Marca	6
			10	00% chance that	at	100 mg	5
				<i>W</i> = 1			
		n - n			2.00		

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X	Y	Ζ	P(X, Y, Z)	X	Y	W	P(X, Y, W)
Coin 1	Coin 2	Bell	<i>Γ</i> (Λ, Ι, Ζ)	Coin 1	Coin 2	witness	P(X, T, VV)
tails	tails	silence	0.25	tails	tails	1	0.125
			,	tails	tails	0	0.125
			E00/ a		change	18th	
			50% (chance	20 2	2
			that	W = 1 that	W = 0		
					2.00		

2019 September 23rd - 27th

Fabio Stella and Luca Bernardinello

Based on Table 2.3 we can easily check that

$P(X = heads | Y = heads) = \frac{0.250}{0.250 + 0.250} = 0.5$

Table 2.3 Probability distribution for two flips of a fair coin and a bell that rings if either flip results in heads, with *X* representing flip one, *Y* representing flip two, and *W* representing a witness who, with variable reliability, reports whether or not the bell has rung.

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X	Y	W	P(X, Y, W)
Coin 1	Coin 2	witness	Γ(Λ, Ι, νν)
heads	heads	1	0.250
heads	tails	1	0.250

$$P(X = heads | Y = heads) = \frac{0.250}{0.250 + 0.250} = 0.5$$

$$P(X = heads) = 0.250 + 0.250 = 0.5$$

$$P(X = heads | Y = heads) = P(X = heads) = 0.5$$

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Table 2.3 Probability distribution for two flips of a fair coin and a bell that rings if either flip results in heads, with *X* representing flip one, *Y* representing flip two, and *W* representing a witness who, with variable reliability, reports whether or not the bell has rung.

X	Y	W	P(X, Y, W)
Coin 1	Coin 2	witness	Ρ(Λ, Τ, Ψ)
heads	heads	1	0.250
heads	tails	1	0.250
tails	heads	1	0.250
tails	tails	1	0.125
tails	tails	0	0.125

In the same way we can show the following

$$P(X = heads | Y = heads) = P(X = heads) = 0.5$$

$$P(X = heads | Y = tails) = P(X = heads) = 0.5$$

$$P(X = tails | Y = heads) = P(X = tails) = 0.5$$

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Therefore, we conclude that X and Y are independent.

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X	Y	W	P(X, Y, W)	
Coin 1	Coin 2	witness	<i>Γ</i> (Λ, Ι, VV)	
heads	heads	1	0.250	
heads	tails	1	0.250	

 $P(X = heads | W = 1) = \frac{0.25 + 0.25}{0.25 + 0.25 + 0.25 + 0.125} = 0.571$

Based on Table 2.3 we can easily check that

Table 2.3 Probability distribution for two flips of a fair coinand a bell that rings if either flip results in heads, with Xrepresenting flip one, Y representing flip two, and Wrepresenting a witness who, with variable reliability, reportswhether or not the bell has rung.

$$P(X = heads | W = 1) = \frac{0.25 + 0.25}{0.25 + 0.25 + 0.25 + 0.125} = 0.571$$

X Coin 1	Y Coin 2	<i>W</i> witness	P(X, Y, W)	$P(X = heads Y = heads, W = 1) = \frac{0.25}{0.25 + 0.25} = 0.5$
heads	heads	1	0.250	
tails	heads	1	0.250	$P(X = heads Y = heads, W = 1) \neq P(X = heads W = 1)$

Therefore, we conclude that X and Y are dependent when conditioning on W = 1.

To summarize

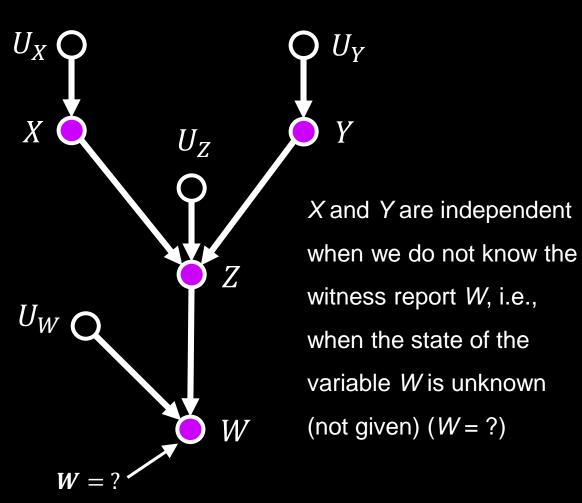
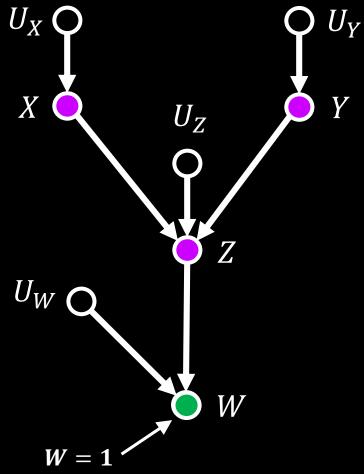
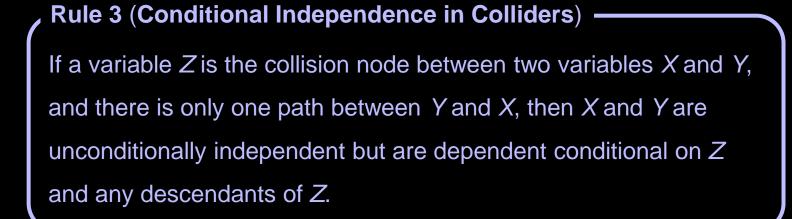


Figure 2.4

X and Y are dependent when we know that the witness reports the bell rings (W = 1), i.e., X and Y become dependent after we know that the witness reports the bell is ringing (W = 1).

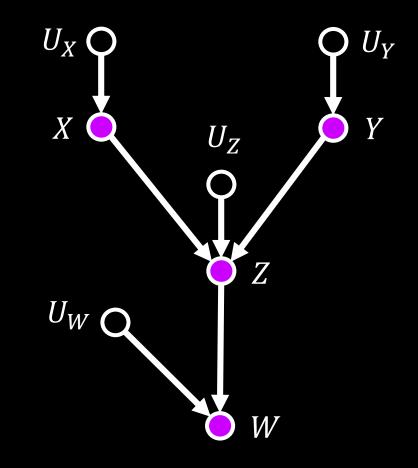


All these considerations lead us to the third rule:



Extremely important to the study of causality, i.e., it allows to:

- test whether a causal model could have generated a data set
- discover models from data
- fully resolve the Simpson's paradox by determining which variables to measure
- estimate causal effect under confounding



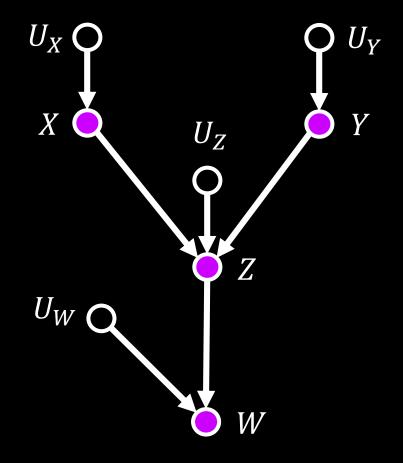


Inquisitive students may wonder why it is that dependencies associated with conditioning on a collider are so surprising to most people as in, for example, the Monty Hall example.

The reason is that humans tend to associate dependence with causation.

Accordingly, they assume (wrongly) that statistical dependence between two variables can only exist if there is a causal mechanism that generates such dependence; that is, either one of the variables causes the other or a third variable causes both.

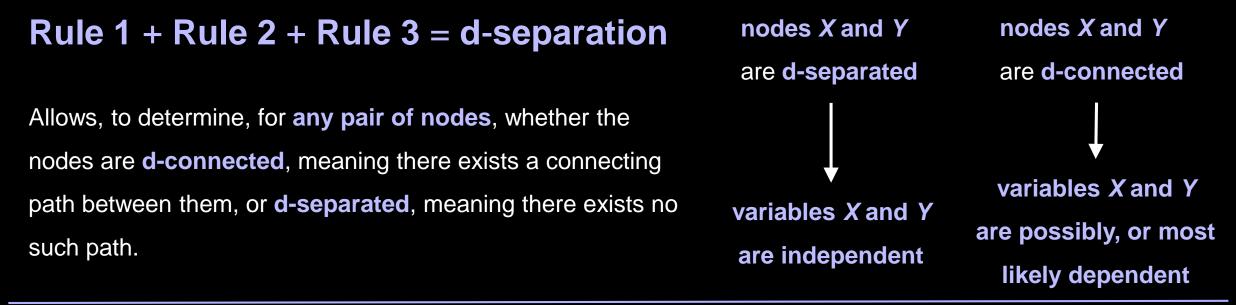
In the case of a collider, they are surprised to find a dependence that is created in a third way, thus **violating the assumption** of "**no correlation without causation.**"



Causal models are generally not as simple as the cases we have examined so far. Specifically, it is rare for a graphical model to consist of a single path between variables.

In most graphical models, pairs of variables will have multiple possible paths connecting them, and each path will traverse a variety of chains, forks, and colliders.

The questions remains whether there is a criterion or process that can be applied to a graphical causal model of any complexity in order to predict dependencies that are shared by all data sets generated by that graph.

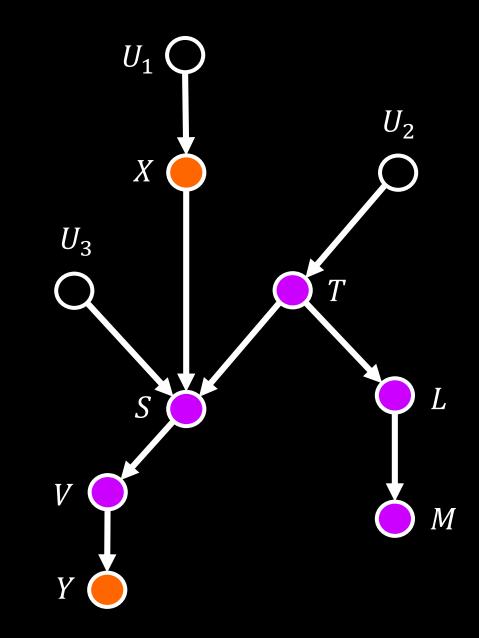


Causal Networks: Learning and Inference

Fabio Stella and Luca Bernardinello

Two nodes *X* and *Y* are **d-separated** if every path between them (should any exist) is **blocked**.

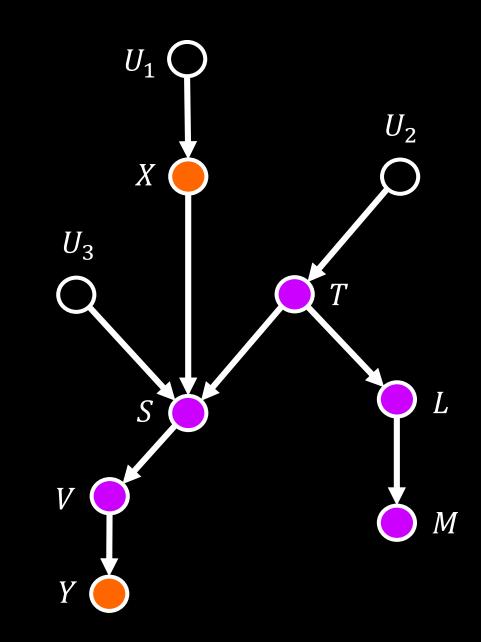
If even one path between *X* and *Y* is **unblocked**, *X* and *Y* are **d-connected**.



Two nodes *X* and *Y* are **d-separated** if every path between them (should any exist) is **blocked**.

If even one path between X and Y is **unblocked**, X and Y are **d-connected**.

• X and Y are d-separated if the following path



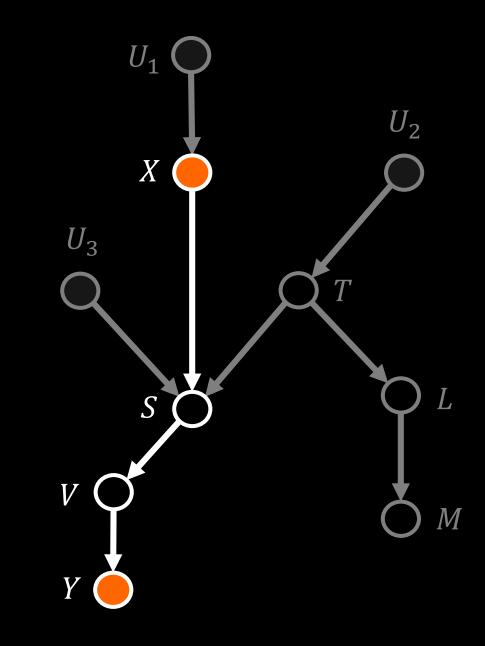
Two nodes *X* and *Y* are **d-separated** if every path between them (should any exist) is **blocked**.

If even one path between X and Y is **unblocked**, X and Y are **d-connected**.

• X and Y are d-separated if the following path

$$X \to S \to V \to Y$$

is blocked.



Two nodes *X* and *Y* are **d-separated** if every path between them (should any exist) is **blocked**.

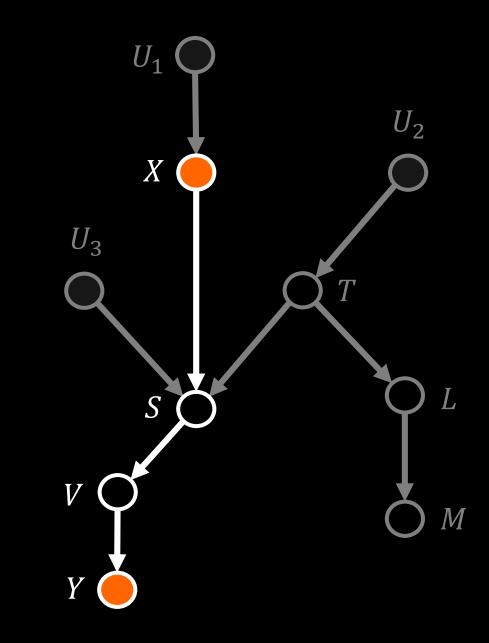
If even one path between X and Y is **unblocked**, X and Y are **d-connected**.

Water flows through Pipes



Path = Pipe

Dependence = Water

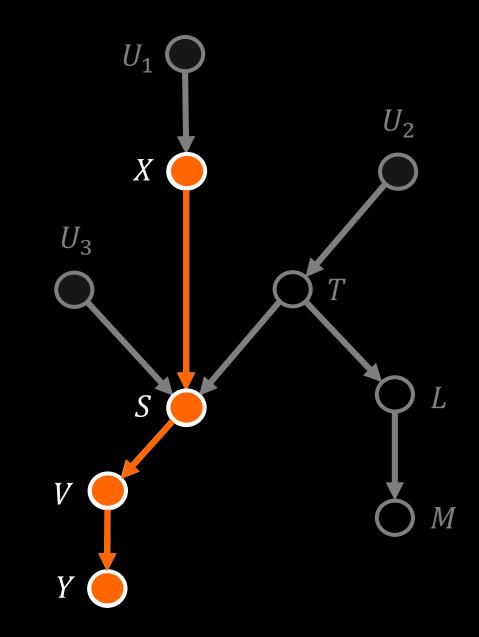


If even one pipe in unblocked, some water can pass from one place to another, and if a single path is clear, the variables at either end will be dependent.

However, a pipe need only be blocked in one place to stop the flow of water through it, it takes only one node to block the passage of dependence in an entire path.

There are certain kinds of nodes that can block a path, depending on whether we are performing unconditional or conditional d-separation.

If we are not conditioning on any variable, then only colliders can block a path.

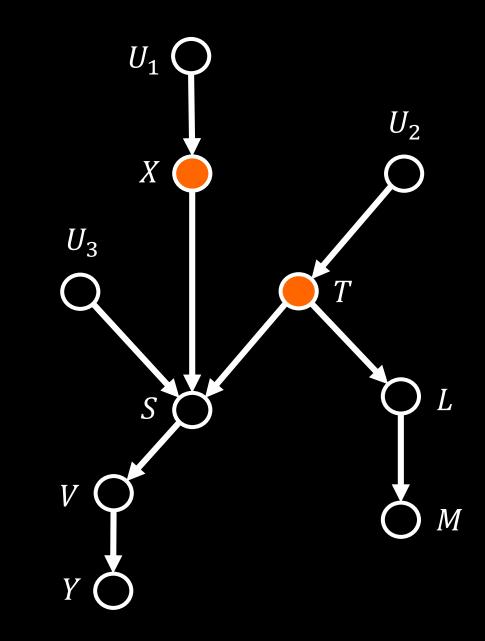


Unconditional d-separation

Consider nodes X and T

There are certain kinds of nodes that can block a path, depending on whether we are performing unconditional or conditional d-separation.

If we are not conditioning on any variable, then only colliders can block a path.



Unconditional d-separation

Consider nodes X and T

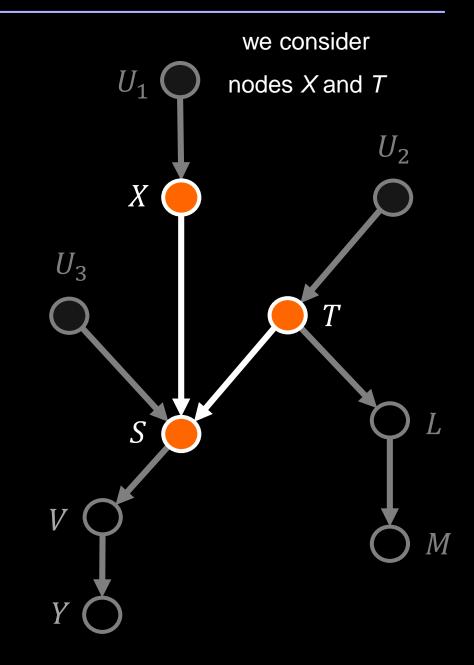
The path

 $X \rightarrow S \leftarrow T$

is blocked by collider S.

There are certain kinds of nodes that can block a path, depending on whether we are performing unconditional or conditional d-separation.

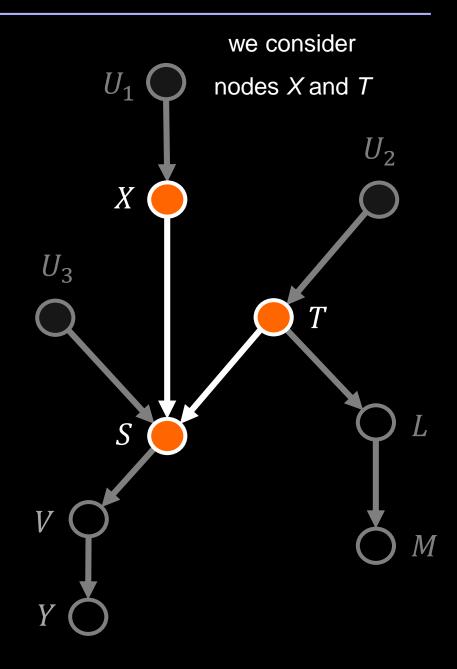
If we are not conditioning on any variable, then only colliders can block a path.



Unconditional d-separation

The reasoning for this is fairly straightforward as we saw in Section 2.3, unconditional dependence can't pass through a collider, i.e. the collider blocks the path.

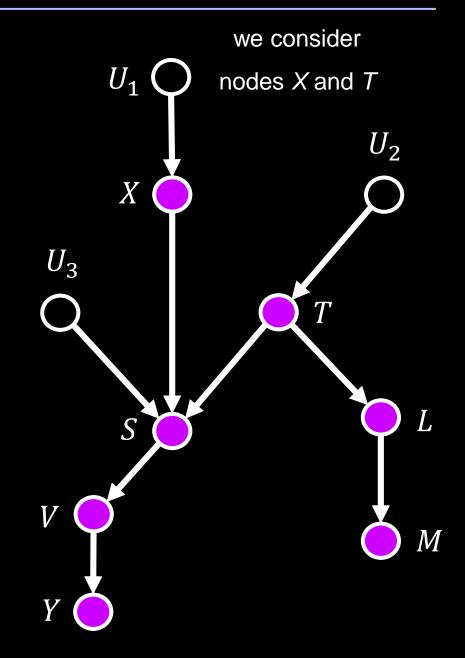
So if every path between two nodes *X* and *Y* has a collider in it, then *X* and *Y* cannot be unconditionally dependent; they must be marginally independent.



Conditional d-separation

If, however, we are conditioning on a set of nodes *Z*, then the following kinds of nodes can block a path:

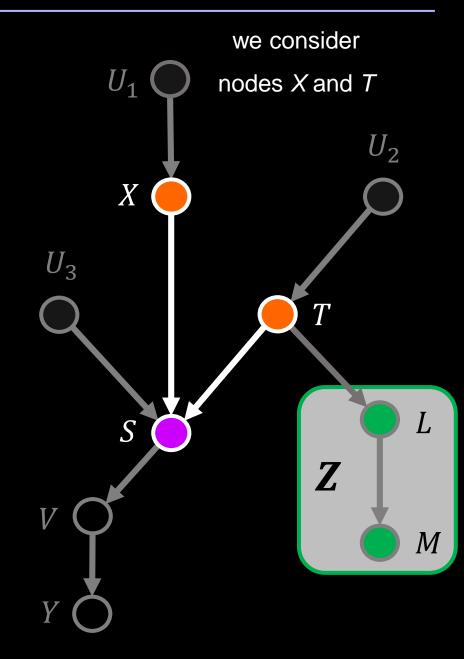
 a collider that is not conditioned on (I.e., not in Z), and that has no descendants in Z.



Conditional d-separation

If, however, we are conditioning on a set of nodes *Z*, then the following kinds of nodes can block a path:

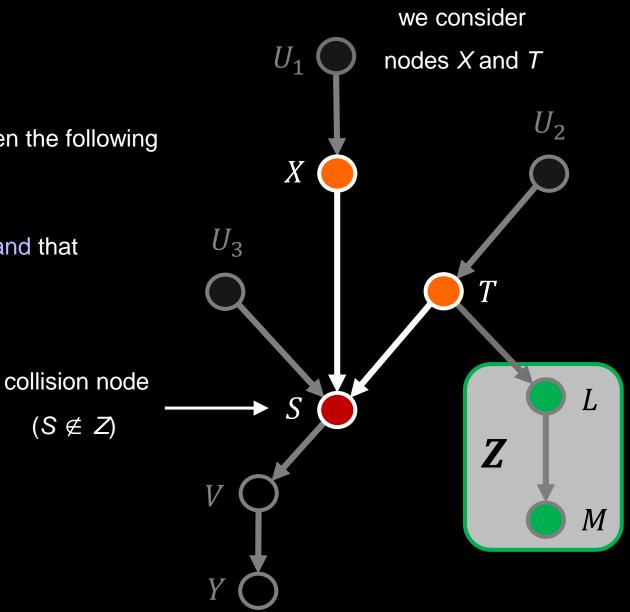
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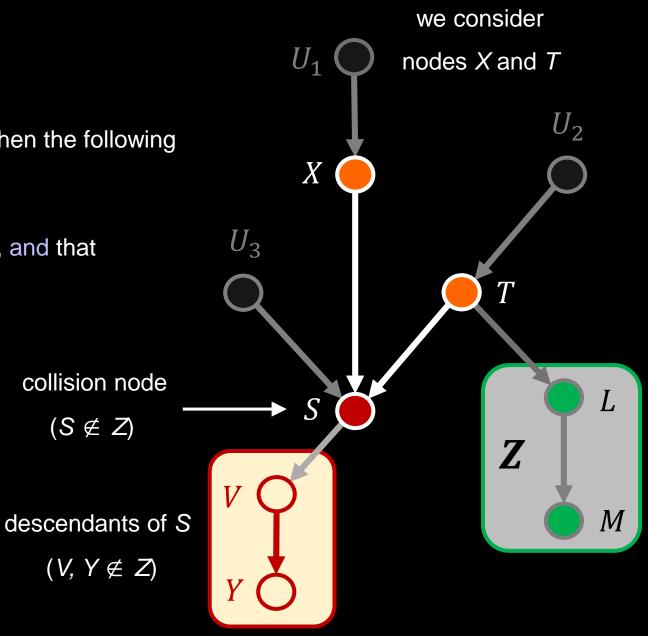


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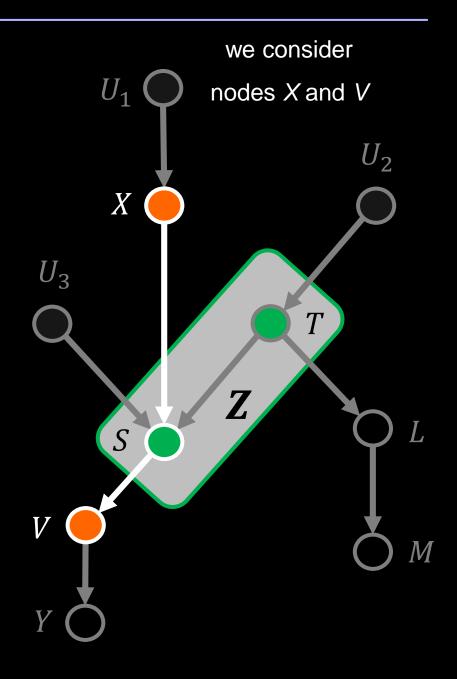
S blocks the path $X \rightarrow S \leftarrow T$



Conditional d-separation

If, however, we are conditioning on a set of nodes Z, then the following kinds of nodes can block a path:

- a collider that is not conditioned on (I.e., not in *Z*), and that has no descendants in *Z*.
- a chain or fork whose middle node is in Z

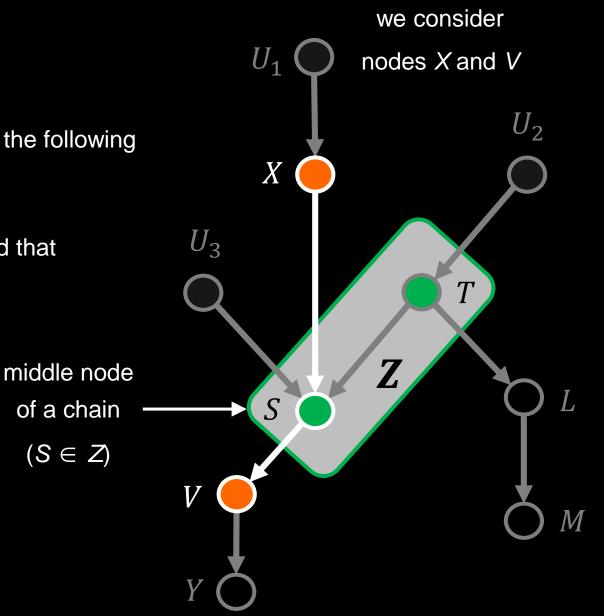


Conditional d-separation

If, however, we are conditioning on a set of nodes *Z*, then the following kinds of nodes can block a path:

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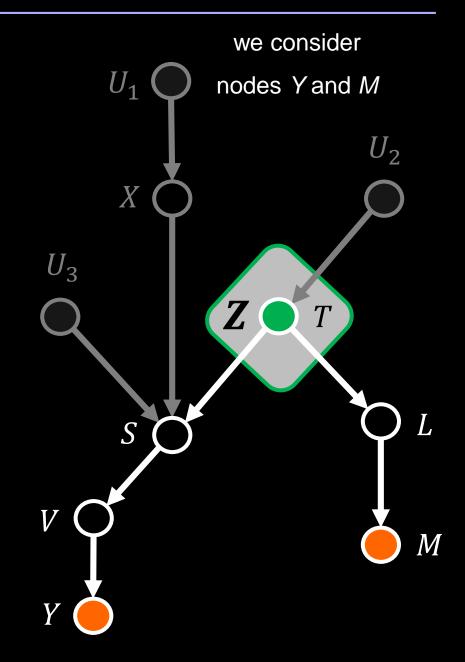
S blocks the path $X \rightarrow S \rightarrow V$



Conditional d-separation

If, however, we are conditioning on a set of nodes *Z*, then the following kinds of nodes can block a path:

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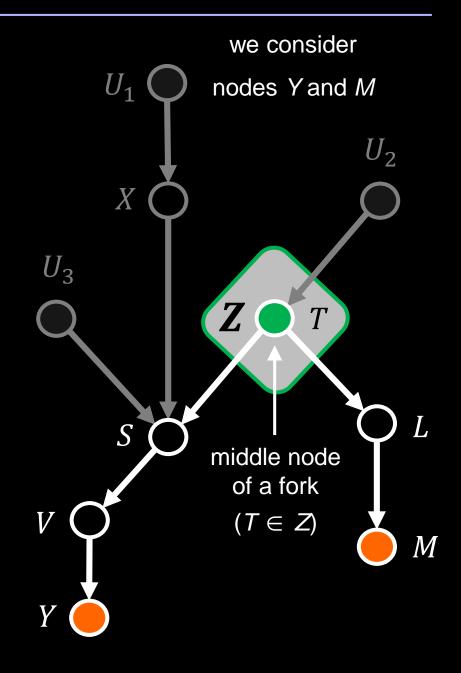


Conditional d-separation

If, however, we are conditioning on a set of nodes *Z*, then the following kinds of nodes can block a path:

- a collider that is not conditioned on (I.e., not in *Z*), and that has no descendants in *Z*.
- a chain or fork whose middle node is in Z

T blocks the path $Y \leftarrow V \leftarrow S \leftarrow T \rightarrow L \rightarrow M$

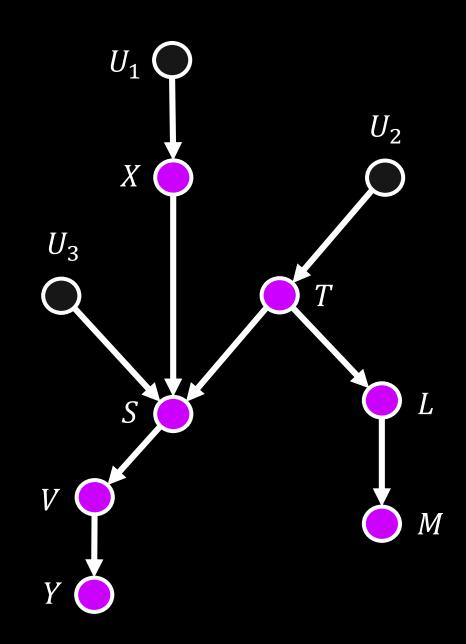


The reasoning behind these points goes back to what we learned in Sections 2.2 and 2.3.

- A collider does not allow dependence to flow between its parents, thus blocking the path,
- but Rule 3 tells us that when we condition on a collider or its descendants, the parent nodes may become dependent.

So

 a collider whose collision node is not in the conditioning set Z would block dependence from passing through a path,



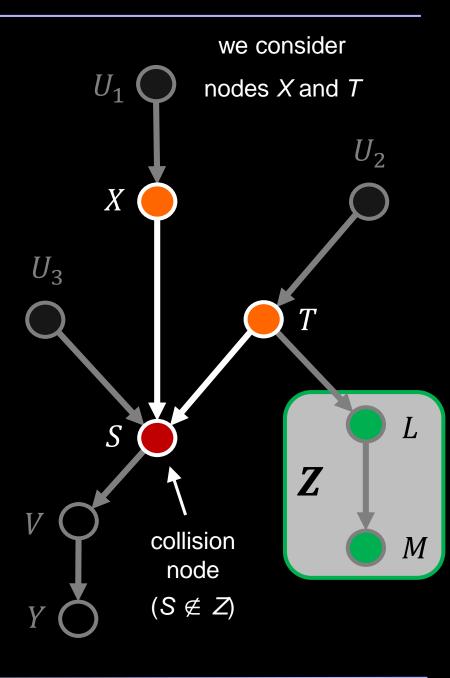
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 a collider whose collision node is not in the conditioning set Z would block dependence from passing through a path,

S blocks the path $X \rightarrow S \leftarrow T$



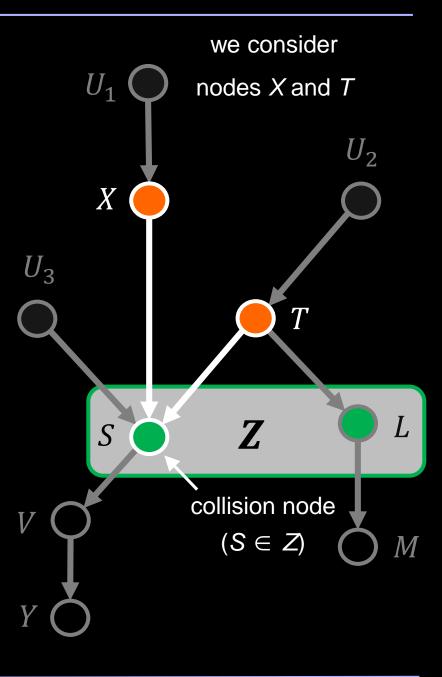
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- A collider does not allow dependence to flow between its parents, thus blocking the path,
- but Rule 3 tells us that when we condition on a collider or its descendants, the parent nodes may become dependent.

So

S does not block (opens) the path $X \rightarrow S \leftarrow T$

 a collider whose collision node or its descendants, is in the conditioning set Z would not block dependence passing through a path.



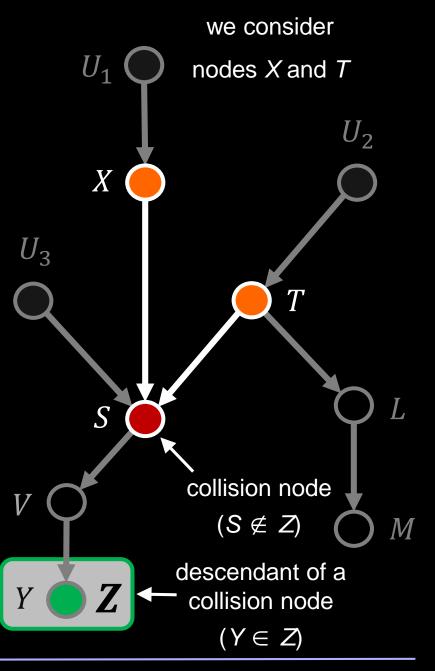
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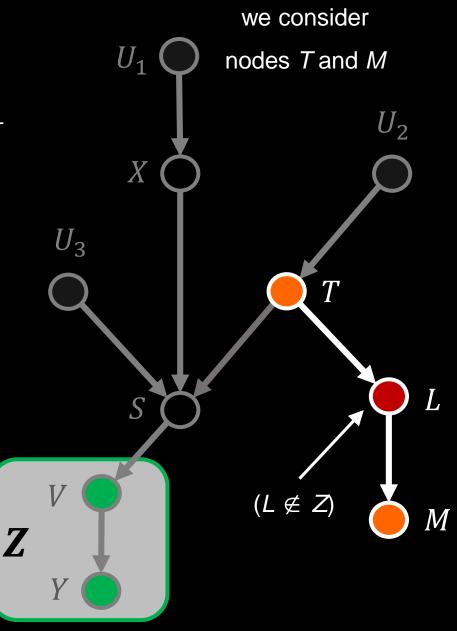
Fabio Stella and Luca Bernardinello

Conversely,

dependence can pass through noncolliders — chains and forks —

L does not block the path $T \rightarrow L \leftarrow M$

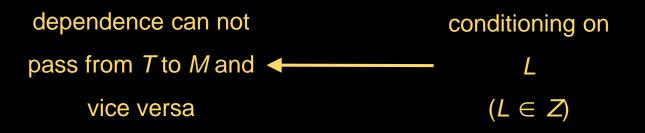
dependence cannot conditioning onpass from T to M andLvice versa $(L \notin Z)$

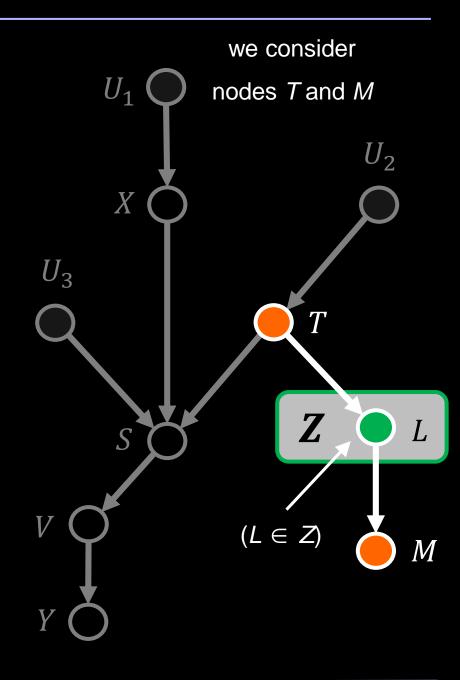


Conversely,

- dependence can pass through noncolliders chains and forks —
- but Rules 1 and 2 tell us that when we condition on them, the variables on either end of those paths become independent (when we consider one path at a time), and thus dependence can not pass through the path.

L blocks the path $T \rightarrow L \leftarrow M$





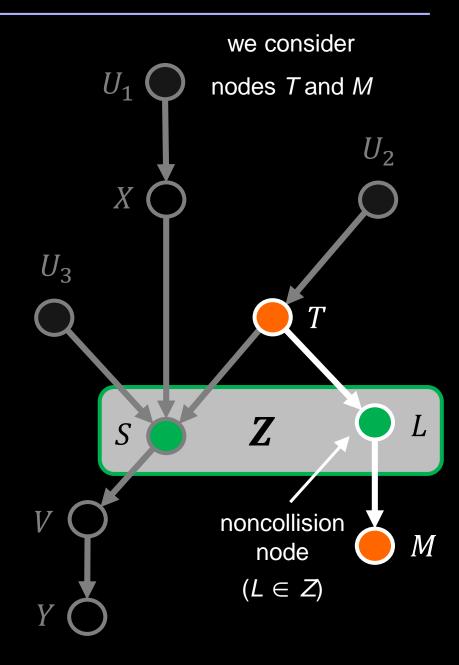
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So

any noncollision node in the conditioning set would block dependence,

L blocks the path $T \rightarrow L \leftarrow M$



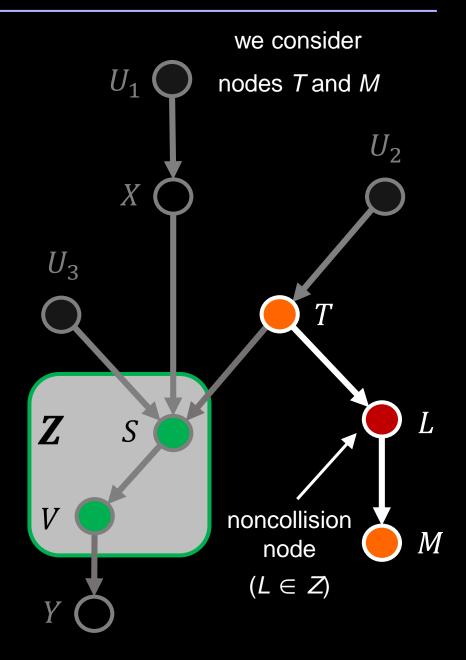
Conversely,

- dependence can pass through noncolliders chains and forks —
- but Rules 1 and 2 tell us that when we condition on them, the variables on either end of those paths become independent (when we consider one path at a time), and thus dependence can not pass through the path.

So

L does not block the path $T \rightarrow L \leftarrow M$

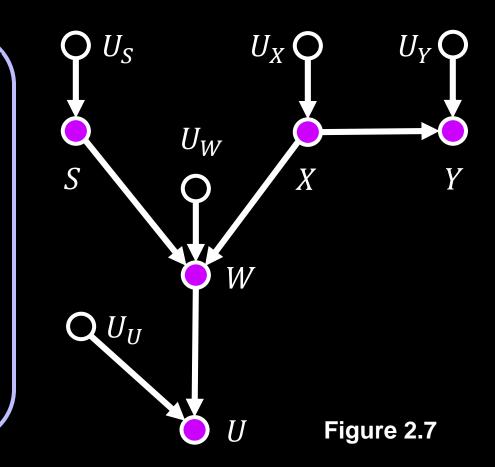
 any noncollision node that is not in the conditioning set would allow dependence through.



Definition 2.4.1 (d-separation) -

A path *p* is blocked by a set of nodes *Z* if and only if

- 1. *p* contains a chain of nodes $A \rightarrow B \rightarrow C$ or a fork $A \leftarrow B \rightarrow C$ such that the middle node *B* is in *Z* (i.e., is conditioned on), or
- 2. *p* contains a collider $A \rightarrow B \leftarrow C$ such that the collision node *B* is not in *Z*, and no descendant of *B* is in *Z*.
- If *Z* blocks every path between two nodes *X* and *Y*, then *X* and *Y* are d-separated, conditional on *Z*, and thus are independent conditional on *Z*.

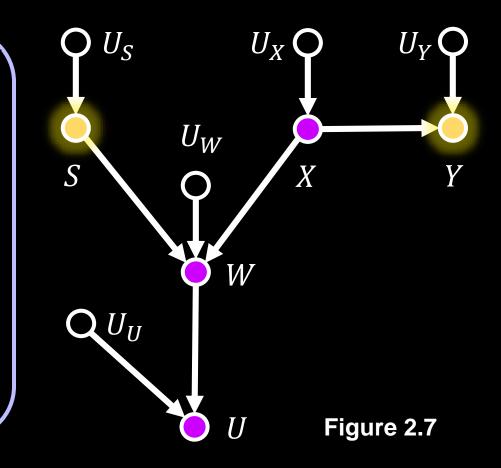


The variables might be discrete, continuous, or a mixture of the two; the relationships between them might be linear, exponential, or any of an infinite number of other relations. No matter the model, however, d-separation will always provide the same set of independencies in the data the model generates.

Definition 2.4.1 (d-separation) -

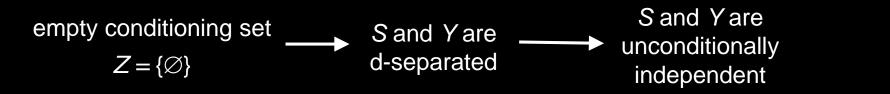
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- If Z blocks every path between two nodes X and Y, then X and Y are d-separated, conditional on Z, and thus are independent conditional on Z.



Why?

In particular, let's look at the relationship between S and Y.



Definition 2.4.1 (d-separation) -

A path *p* is blocked by a set of nodes *Z* if and only if

- 1. *p* contains a chain of nodes $A \rightarrow B \rightarrow C$ or a fork $A \leftarrow B \rightarrow C$ such that the middle node *B* is in *Z* (i.e., is conditioned on), or
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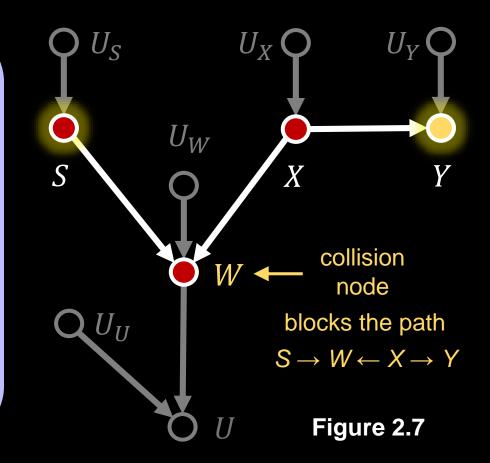
 U_X C OU_S U_{Y} U_W S W OU_{II} Figure 2.7

Because there is no unblocked path between S and Y.

Definition 2.4.1 (d-separation) -

A path *p* is blocked by a set of nodes *Z* if and only if

- 1. *p* contains a chain of nodes $A \rightarrow B \rightarrow C$ or a fork $A \leftarrow B \rightarrow C$ such that the middle node *B* is in *Z* (i.e., is conditioned on), or
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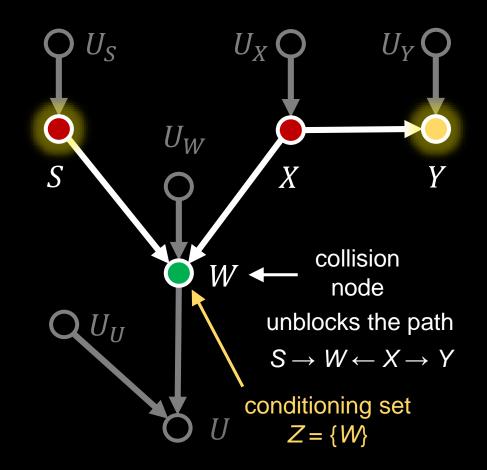
Because there is no unblocked path between *S* and *Y*.

There is only one path between *S* and *Y*, and that path is blocked by a collider ($S \rightarrow W \leftarrow X$).

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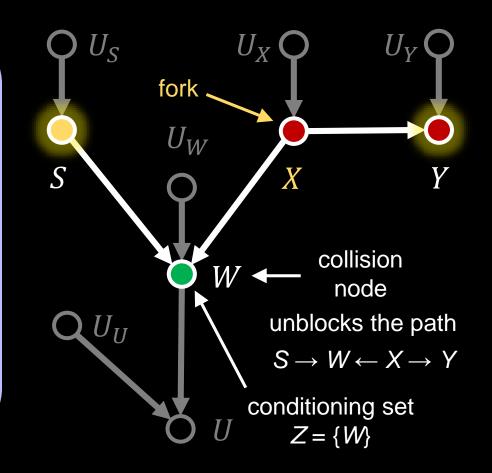
But suppose we condition on *W*.

d-separation tells us that S and Y are d-connected, conditional on W.

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- If Z blocks every path between two nodes X and Y, then X and Y are d-separated, conditional on Z, and thus are independent conditional on Z.



The reason is that our conditioning set is now $Z = \{W\}$, and since the only path between S and Y contains a fork (X) that is not in that set, and the only collider (W) on the path is in that set, that path is not blocked.

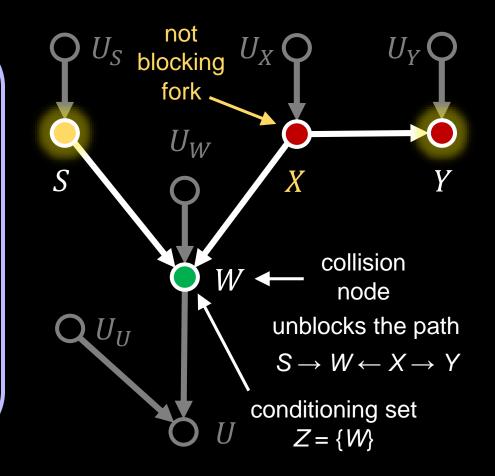
(Remember that conditioning on colliders "unblocks" them.)

2019 September 23rd - 27th

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2019 September 23rd - 27th

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descendant of a collision node U_U $U \leftarrow Conditioning set$ $U \leftarrow Z = \{U\}$ $C \leftarrow X \rightarrow Y$

 U_{S}

S

fork

 U_W

 U_X

 U_{Y} (

collision

node

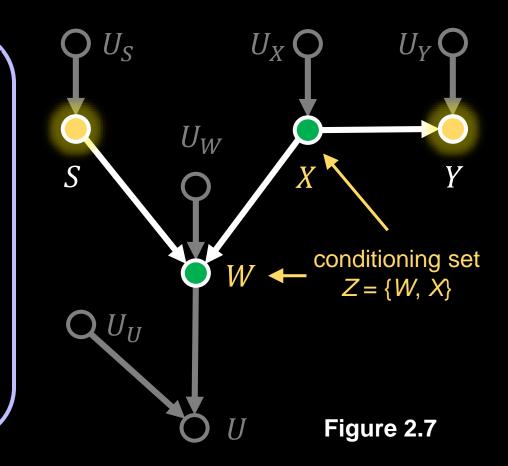
The same is true if we condition on U, because U is a

descendant of a collider along the path between S and Y.

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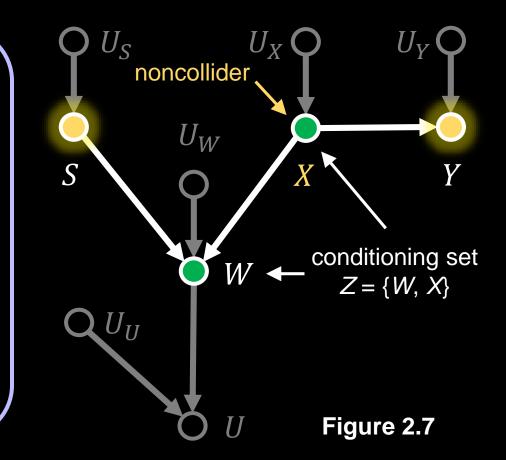
On the other hand, if we condition on the set $Z = \{W, X\}$, S and Y remain independent.

This time, the path between S and Y is blocked by the first criterion, rather than the second.

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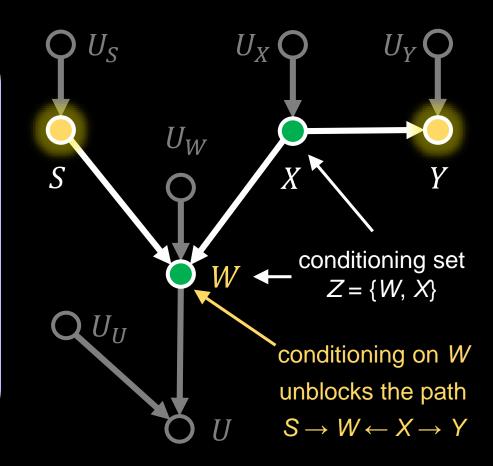
There is now a noncollider node (X) on the path that is in the conditioning set.

Fabio Stella and Luca Bernardinello

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Though

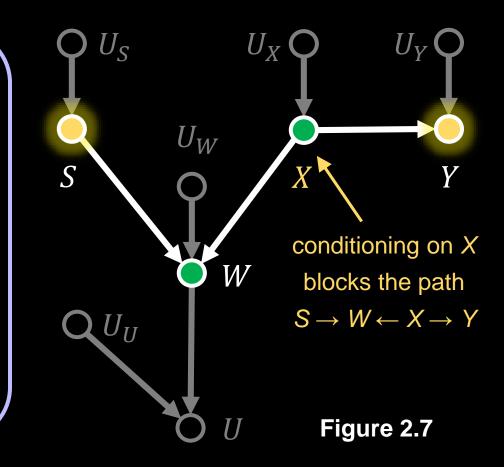
$$S \to W \leftarrow X \to Y$$

has been unblocked by conditioning on *W*, one blocked node is sufficient to block the entire path.

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Since the only path between *S* and *Y* is blocked by this conditioning set *Z*, *S* and *Y* are d-separated conditional on $Z = \{W, X\}$.

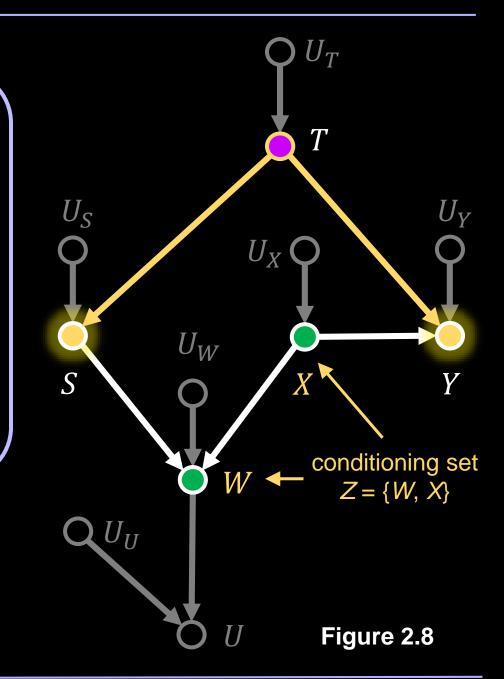
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Now, consider what happens when we add another path between *S* and *Y*, as in Figure 2.8. Why?

S and Y are now unconditionally dependent.

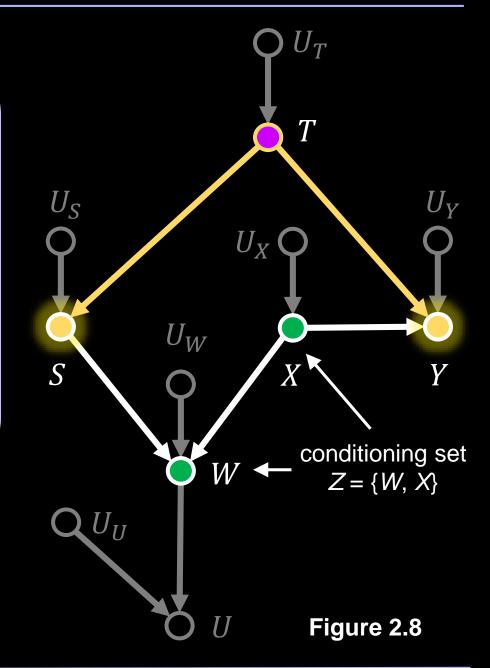


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Because there is a path between them $(S \leftarrow T \rightarrow Y)$ that contains no colliders, and the middle node *T* does not belong to the conditioning set $Z = \{W, X\}$.

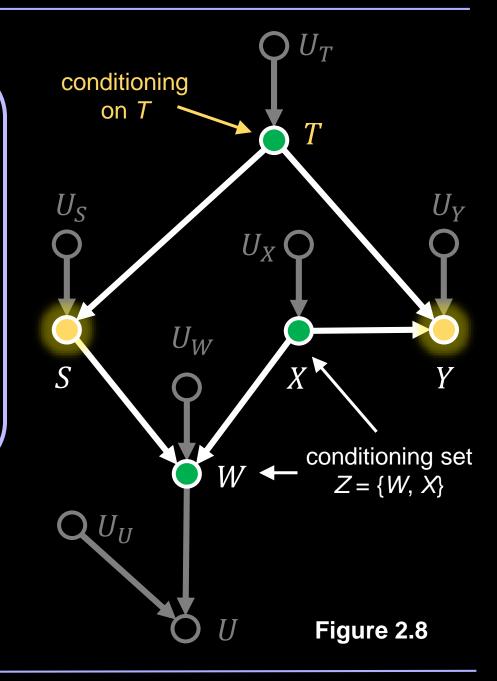


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- If *Z* blocks every path between two nodes *X* and *Y*, then *X* and *Y* are d-separated, conditional on *Z*, and thus are independent conditional on *Z*.

If we also condition on *T*, *i.e.*, *if we set* $Z = \{W, X, T\}$, however, the path ($S \leftarrow T \rightarrow Y$) is blocked, and *S* and *Y* become independent again.



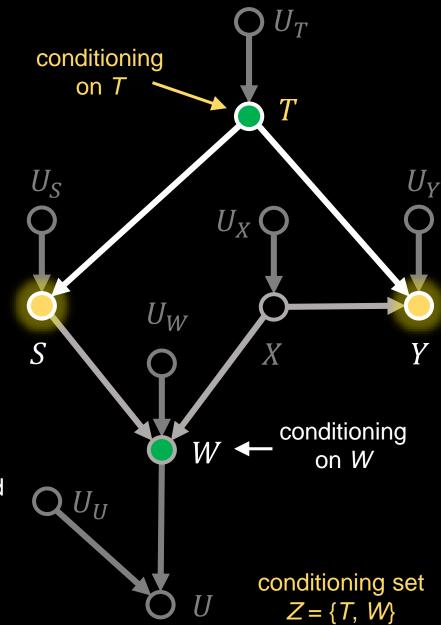
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Conditioning on $Z = \{T, W\}$, on the other hand, makes them d-connected again:

• conditioning on *T* blocks the path $S \leftarrow T \rightarrow Y$,



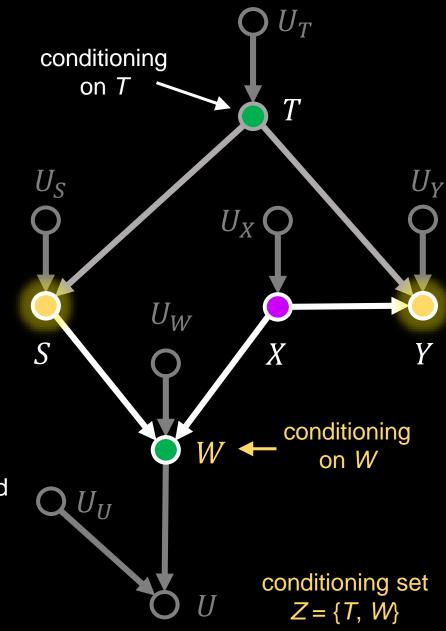
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Conditioning on $Z = \{T, W\}$, on the other hand, makes them d-connected again:

- conditioning on T blocks the path $S \leftarrow T \rightarrow Y$,
- but conditioning on W unblocks the path $S \rightarrow W \leftarrow X \rightarrow Y$.



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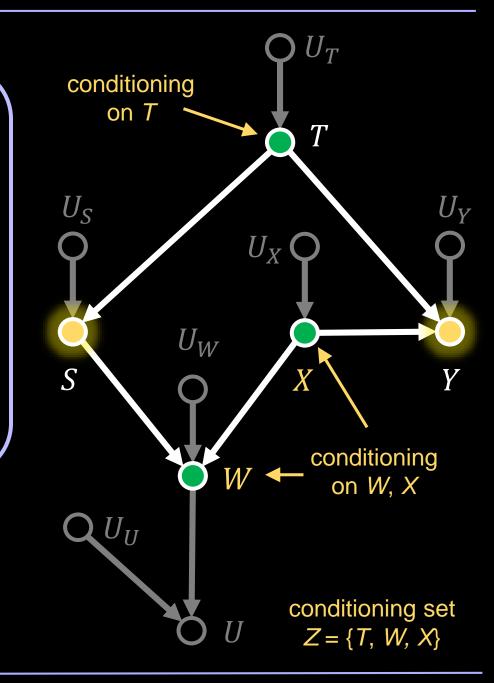
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And if we add X to the conditioning set, making it

 $Z = \{T, W, X\},\$

S, and Y become independent yet again!



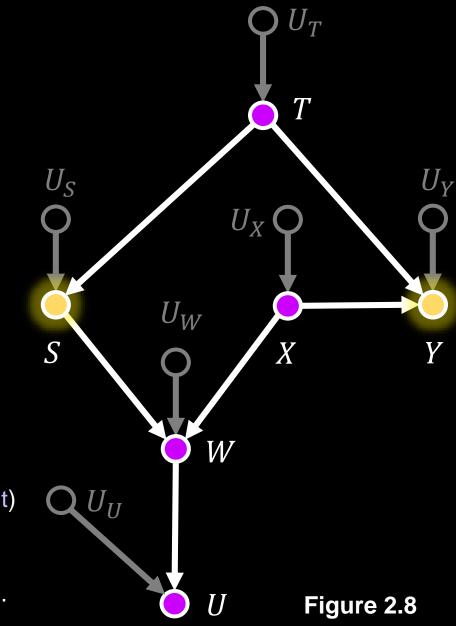
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In this graph, S and Y are d-connected (and therefore likely dependent) conditional on

W, *U*, {*W*, *U*}, {*W*, *T*}, {*U*, *T*}, {*W*, *U*, *T*}, {*W*, *X*}, {*U*, *X*}, and {*W*, *U*, *X*}.

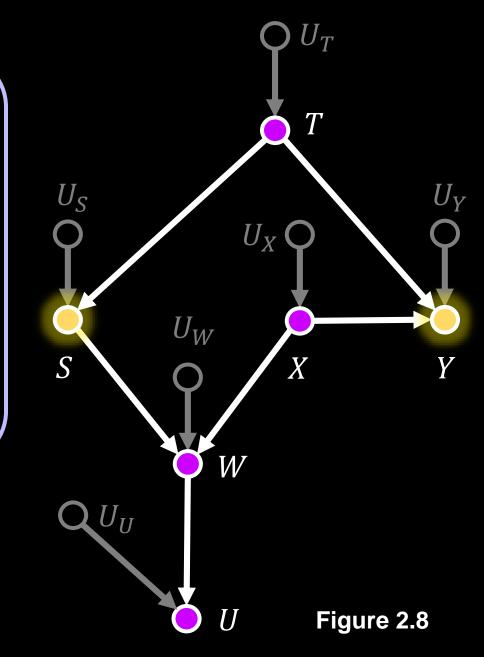


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S and Y are d-separated (and therefore independent) conditional on: T, {X, T}, {W, X, T}, {U, X, T}, and {W, U, X, T}.



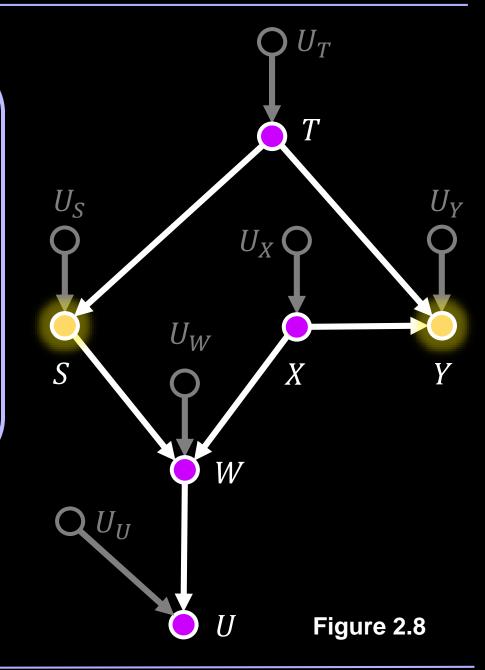
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Note that T is in every conditioning set that d-separates S and Y.

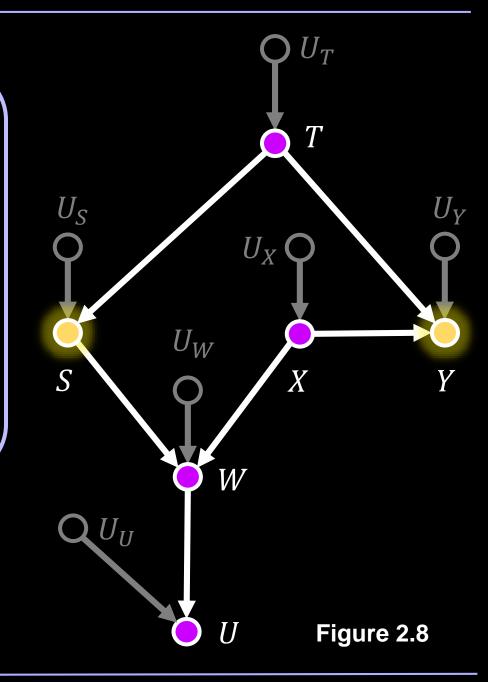


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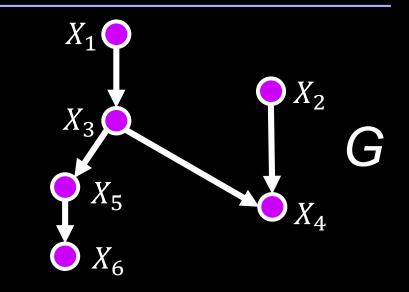
T is in every conditioning set that d-separates *S* and *Y* because *T* is the only node in a path that unconditionally d-connects *S* and *Y*, so unless it is conditioned on, *S* and *Y* will always be d-connected.



The preceding sections demonstrate that causal models have testable implications in the data sets they generate.

For instance, if we have a graph *G* that we believe might have generated a data set S, d-separation will tell us which variables in *G* must be independent conditional on which other variables.

Conditional independence is something we can test for using a data set S.



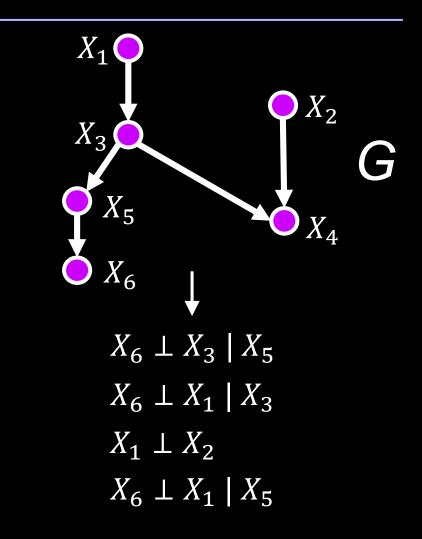
	N/	N/	N/	N/	14
Χ1	X 2	Х 3	X 4	X 5	Х ₆
0	1	1	1	0	1
0	1	0	1	1	1
1	0	1	1	0	0
0	0	1	0	1	1
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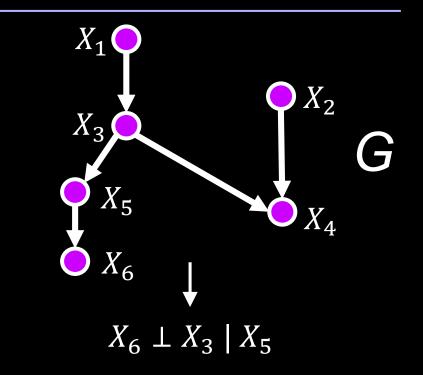


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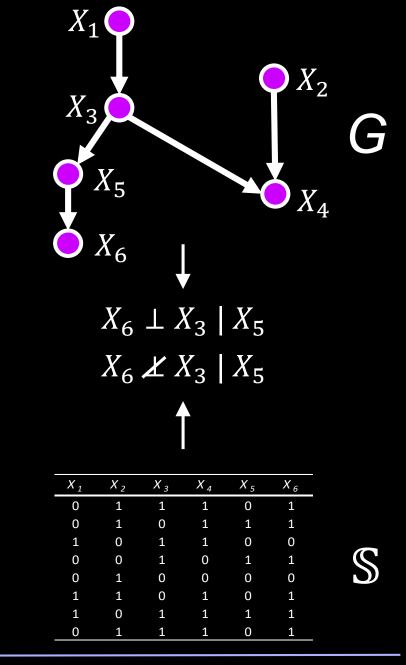


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- Suppose we list the d-separation conditions in *G*, and note that variables X_6 and X_3 must be independent conditional on X_5 .
- Then, suppose we estimate the probabilities based on S, and discover that the data suggests that X_6 and X_3 are not independent conditional on X_5

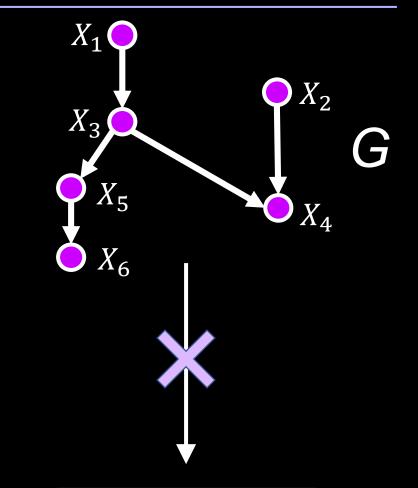


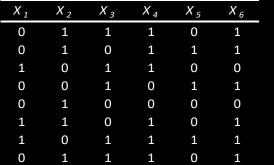
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- Then, suppose we estimate the probabilities based on S, and discover that the data suggests that X_6 and X_3 are not independent conditional on X_5
- We can then reject G as a possible causal model for S.





We can demonstrate it on the causal model of Figure 2.9.

Among the many conditional independencies advertised by the model, we find that W and Z_1 are independent given X, because X d-separates W from Z_1 .

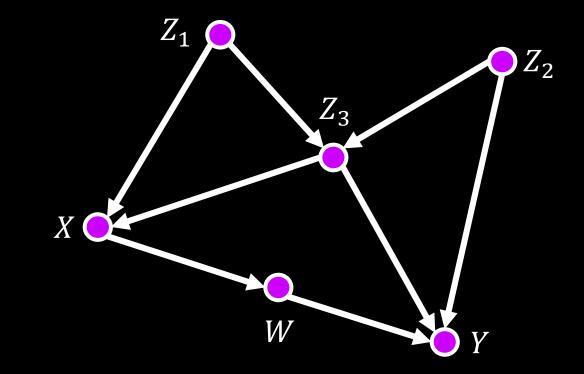
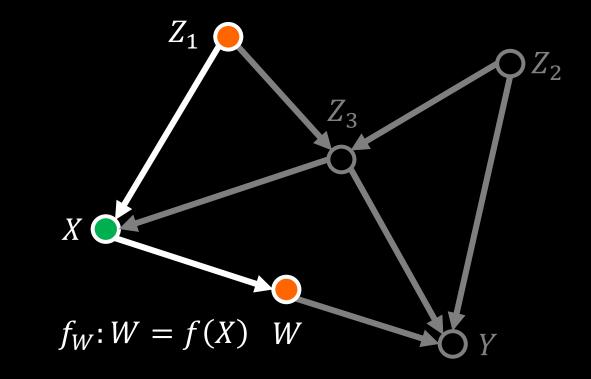


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Now suppose we regress W on X and Z_1 . Namely, we find the line

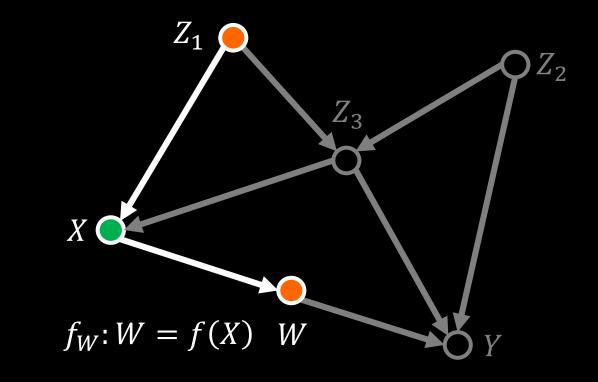
$$W = r_X X + r_1 Z_1$$

that best fits our data.

IF $r_1 \neq 0 \Rightarrow W$ depends on Z_1 given X

and consequently, that the model in Figure 2.9 is wrong.

[Conditional correlation implies conditional dependence.]





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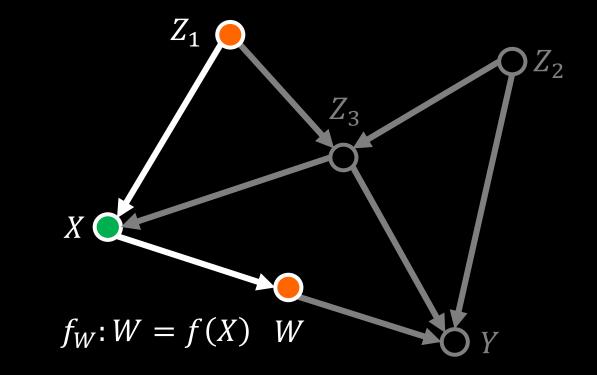


Figure 2.9

Not only do we know that the model in **Figure 2.9** is wrong, but we also know where it is wrong;

 the true model must have a path between W and Z₁ that is not d-separated by X.

Finally, this is a theoretical result that holds for all acyclic models with independent errors (*Verma and Pearl 1990*), and we also know that if every d-separation condition in the model matches a conditional independence in the data, then no further test can refute the model.

This means that, for any data set whatsoever, one can always find a set of functions *F* for the model and an assignment of probabilities to the *U* terms, so as to generate the data precisely.

There are other methods for testing the fitness of a model.

The standard way of evaluating fitness involves a **statistical hypothesis test** over the entire model, that is, we evaluate how likely it is for the observed samples to have been generated by the hypothesized model, as opposed to sheer chance.

However, since the model is not fully specified, we need to first estimate its parameters before evaluating that likelihood. This can be done (approximately) when we assume a linear and Gaussian model (i.e., all functions in the model are linear and all error terms are normally distributed), because, under such assumptions, the joint distribution (also Gaussian) can be expressed succinctly in terms of the model's parameters, and we can then evaluate the likelihood that the observed samples to have been generated by the fully parameterized model (Bollen 1989).

There are, however, a number of issues with this procedure:

- if any parameter cannot be estimated, then the joint distribution cannot be estimated, and the model cannot be tested. (this can occur when some of the error terms are correlated or, equivalently, when some of the variables are unobserved)
- this procedure tests models globally. If we discover that the model is not a good fit to the data, there is no way for us to determine why that is—which edges should be removed or added to improve the fit.
- when we test a model globally, the number of variables involved may be large, and if there is
 measurement noise and/or sampling variation associated with each variable, the test will not be reliable.

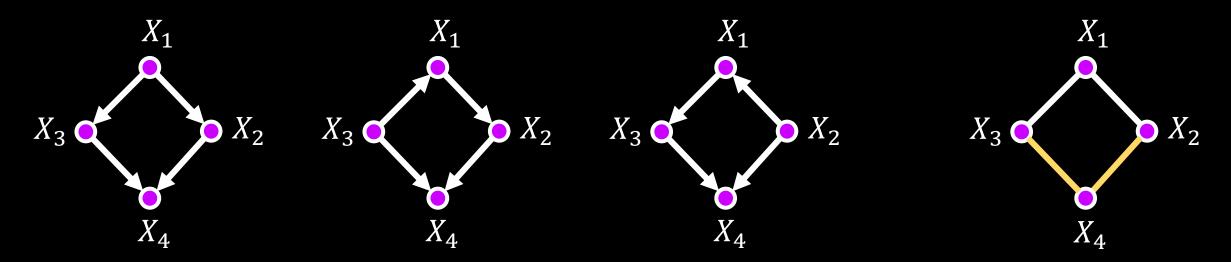
d-separation presents several advantages over this global testing method.

- it is nonparametric, meaning that it doesn't rely on the specific functions that connect variables; instead, it uses only the graph of the model in question,
- it tests models locally, rather than globally. This allows us to identify specific areas, where our hypothesized model is flawed, and to repair them, rather than starting from scratch on a whole new model. It also means that if, for whatever reason, we can't identify the coefficient in one area of the model, we can still get some incomplete information about the rest of the model. (As opposed to the first method, in which if we could not estimate one coefficient, we could not test any part of the model.)

If we had a computer, we could test and reject many possible models in this way, eventually whittling down the set of possible models to only a few whose testable implications do not contradict the dependencies present in the data set. It is a set of models, rather than a single model, because some graphs have indistinguishable implications. A set of graphs with indistinguishable implications is called an **equivalence class**.

Two graphs G_1 and G_2 are in the same equivalence class if they share a common skeleton—that is,

- the same edges, regardless of the direction of those edges—and
- if they share common v-structures, that is, colliders whose parents are not adjacent.

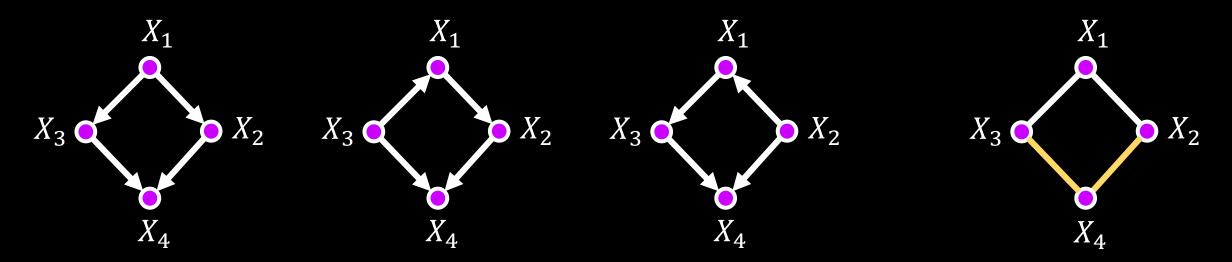


Three equivalent graphs and their skeleton with the common v-structure highlighted.

Any two graphs that satisfy this criterion have identical sets of d-separation conditions and, therefore, identical sets of testable implications (Verma and Pearl 1990).

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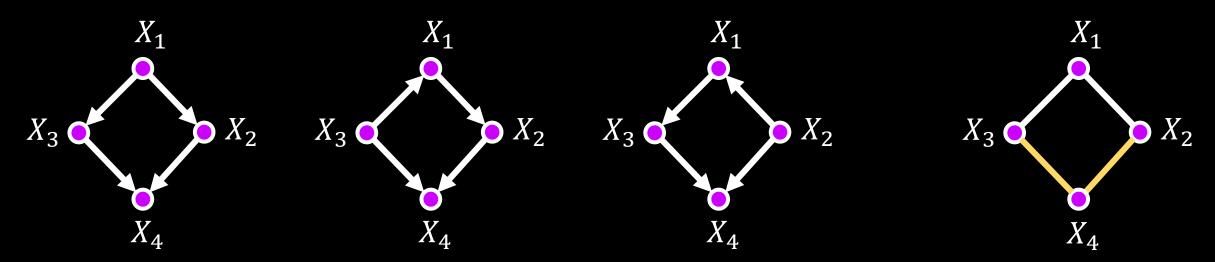


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The importance of this result is that it allows us to search a data set for the causal models that could have generated it. Thus, not only can we start with a causal model and generate a data set—but we can also start with a data set, and reason back to a causal model.

This is enormously useful, since the object of most data-driven research is exactly to find a model that explains the data.



Three equivalent graphs and their skeleton with the common v-structure highlighted.