

MICROECONOMETRIA - LEZIONE 7

~~MODELLO PANEL EFFICIENTI RANDOM~~ / Error component model

$$y_{it} = x_{it}\beta + u_{it}, \text{ dove } u_{it} = \mu_i + v_{it}$$

T=1 T=K T=1

(Errorne composite)

$\hat{\beta}_{GLS} \leftarrow$ EFFICIENZA INDIVIDUALE
VAI COSTANTE CON X_{it}

$$\Downarrow$$

$$E(y_i | x_i) \equiv V = \sigma_y^2 C_T + \sigma_w^2 D_T$$

FIT

$$\sigma_y^2 \equiv T \sigma_\mu^2 + \sigma_w^2$$

Duke

$$GLS \xrightarrow{\text{STINA 21}} FGLS$$

$$\sigma_w^2 E \sigma_w^2$$

STINATA DI σ^2

SI SFRUTTA LA MASSIMIZZAZIONE BETWEEN

(TRANSFORMAZIONE
KI NOBIS DISPERTE A T)

MODELLO RANDOM EFFECTS (FORMA SCALARE)

$$(1) y_{it} = \alpha + \sum_{k=2}^K \beta_k x_{kit} + \underbrace{(\mu_i + u_{it})}_{v_{it}}$$

Transformations y_{it} , x_{2it} ($R=2-k$), μ_i , U_{it}

can BETWEEN :

$$(2) y_{it} = \alpha + \sum_{r=2}^k \beta_r x_{rit} + \underbrace{(\mu_i + U_{it})}_{\bar{u}_i}$$

DIVE $z_{it} = \frac{1}{T} \sum_t z_{it}$
→ cross-section ($i = 1 \sim N$)

$$(3) \text{VAR}(N_{i.}) = \text{VAR}(\mu_{i.} + U_{i.}) = \text{VAR}(\mu_{i.}) + \text{VAR}(U_{i.})$$

$$= \tilde{\sigma}_{\mu}^2 + \text{VAR}\left(\frac{1}{T} \sum_{t=1}^T U_{it}\right) \quad \left(\begin{array}{l} \mu_{i.} \in U_{it} \cdot X_{it} \\ \text{CAMBANGAN PEN} \\ \text{POTRESI} \end{array} \right)$$

$$= \tilde{\sigma}_{\mu}^2 + \frac{1}{T^2} \text{VAR}\left(\sum_{t=1}^T U_{it}\right)$$

$$= \tilde{\sigma}_{\mu}^2 + \frac{1}{T^2} \cdot T \cdot \tilde{\sigma}_w^2 \quad \downarrow$$

$$T \text{VAR}(n_{i.}) = T \sigma_{\mu}^2 + \sigma_w^2$$

Calcolazione: STIMARE $\text{VAR}(n_{i.})$ E MOLTIPLICARLO

PER T SI OTTIENE LA STING DI σ_{μ}^2

PER STIMARE $\text{VAR}(n_{i.})$ USIAMO LA SECONDA PROCEDURA:

1) STIMARE CON OLS IL MODELLO (2), OTTENERE I

$$\text{RESIDU: } \hat{v}_i = y_i - \sum_{k=2}^K \hat{\beta}_k x_{ki} - \hat{\alpha}$$

2) SULLA BASE DEI \hat{v}_i OTTENERE LA VARIANZA
STIMATA $\widehat{\text{var}}(\hat{v}_i)$

$$3) \hat{\sigma}_v^2 = \frac{1}{n} \widehat{\text{var}}(\hat{v}_i)$$

$$\Downarrow$$
$$\hat{V} = \hat{\sigma}_T^2 C_T + \hat{\sigma}_w^2 D_T \Rightarrow \underline{\underline{FGLS}}$$

MINIMUM VARIANCE ESTIMATION:

$$\hat{\sigma}_y^2 = \frac{\hat{\sigma}_T^2 - \hat{\sigma}_w^2}{T}$$

STINGA DEGLI EFFETTI INDIVIDUALI RANDOM ($\hat{\mu}_i$)

PREVISIBILI CON MODELLI DI REGRESSIONE LINEARI

$N = 1$ (NO DATI PANEL)

$t = 1 \dots T$ (OMITTORE DI STINGA)

$t = T+1 \dots T+G$ (OMITTORE DI PREVISIBILE)

MODELLI DI REGRESSIONE CLASSICI

$$y = X\beta + U$$

$T \cdot 1$ $T \cdot K$ $K \cdot 1$ $T \cdot 1$

$$T = 1 \dots T$$

$$E(UU') = I_T \quad (\sigma_u^2 = 1)$$

$$y_F = X_F\beta + U_F$$

$G \cdot 1$ $G \cdot K$ $K \cdot 1$ $G \cdot 1$

$$T = T + 1 \dots T + G$$

$$E(U_F U_F') = I_G$$

STABILITÀ
DEI PARAMETRI (CON IPOWER)

$$E(UU') = 0$$

Prueviasi : $\hat{y}_F = X_F \hat{\beta}$

$$G \cdot 1 \quad G \cdot k \quad k \cdot 1, \text{ dove } \hat{\beta}_{k \cdot 1} = (X \cdot X)^{-1} X \cdot 1 y$$

← VALORI "PUNTI" DI X SONO NOTI

← SINGOLITÀ DI $\hat{\beta}$

Modello di regressione GLS

$$y = X\beta + u, \quad E(uu') = \Omega \quad t=1 \dots T$$

$$\hookrightarrow \hat{\beta}_{GLS} = (X' \Omega^{-1} X)^{-1} X' \Omega^{-1} y \quad \uparrow \text{"GLS"}$$

$$y = X\beta + u, \quad E(uu') = \Omega \quad (t=1 \dots T+G)$$

$G:1 \quad G:k \quad k:1 \quad G:1$ DANE $E(uu') \neq 0$ $\Omega \neq 0$

PREVISIJA: $\hat{Y}_F = X_F \hat{\beta}_{OLS} + Q' \Omega^{-1} U$

$\underbrace{\quad}_{G \cdot 1}$
 $\underbrace{\quad}_{G \cdot K}$
 $\underbrace{\quad}_{K \cdot 1}$
 $\underbrace{\quad}_{G \cdot T}$
 $\underbrace{\quad}_{T \cdot T}$
 $\underbrace{\quad}_{T \cdot 1}$

PREVISIJA $G \cdot 1$
 \downarrow
 $t = 1 \dots T$

MODEL PANEL EFECTI RANDOM ($i = 1 \dots N$)

$y_i = \alpha_i \delta + \nu_i$, $\nu_i = \mu_i \gamma + u_i$ ($t = 1 \dots T$)
 $T \cdot 1$ $T \cdot K$ $K \cdot 1$ $T \cdot 1$ $T \cdot 1$ $1 \cdot 1$ $T \cdot 1$ $T \cdot 1$

POSTSI: $G = 1$ (PREVISIJA IZ PREVIŠTA)

$$y_{i,T+1} = x_{i,T+1} \delta + \underbrace{\varepsilon_{i,T+1}}_{(1,1)}$$

$y_{i,T+1}$
(scalar)

PREVIOUS:

$$\hat{y}_{i,T+1} = x_{i,T+1} \hat{\delta}_{OLS} + Q' V^{-1} \hat{\varepsilon}_i$$

(1,1) (1,K) (K,1) (1,T) (T,T) T,1

RESIDUALS
RANDOM

$$\text{DVE } \hat{\varepsilon}_i = y_i - x_i \hat{\delta}_{OLS}$$

$$\begin{aligned}
 E Q &= E \left(\underbrace{\begin{matrix} \mu_i & \mu_{i+1} \\ \sigma_i^2 & \sigma_{i+1}^2 \end{matrix}}_{(T-1)} \right) \\
 &= E \left[\begin{matrix} \mu_i & \mu_{i+1} \\ \sigma_i^2 & \sigma_{i+1}^2 \end{matrix} \right] \\
 &= E(\mu_i^2) = \sigma_i^2
 \end{aligned}$$

$$V = \sigma_1^2 G + \sigma_w^2 D \Rightarrow V^{-1} = \frac{1}{\sigma_1^2} G + \frac{1}{\sigma_w^2} D$$



$$\hat{y}_{i+1} = x_{i+1} \hat{\delta}_{OLS} + (\sigma_u^2 J)' \left(\frac{1}{\sigma_1^2} G + \frac{1}{\sigma_w^2} D \right) \hat{v}_i$$

$$= x_{i+1} \hat{\delta}_{OLS} + \left[\frac{\sigma_u^2}{\sigma_1^2} \begin{pmatrix} J' G \\ J' \end{pmatrix} + \frac{\sigma_u^2}{\sigma_w^2} \begin{pmatrix} J' D \\ J' \end{pmatrix} \right] \hat{v}_i$$

N.B. $J' G = J' \cdot J J' = \cancel{J' J} \cdot J' = J'$

$$J'D_T = J'(I_T - C_T) = J' - J'C_T = J' - J' = 0$$

↓

$$\hat{y}_{i,t+1} = x_{i,t+1} \hat{\delta}_{OLS} + \frac{\hat{\sigma}_u^2}{\sigma_u^2}$$

(i.1) $1 \times k$ $k \times 1$ 1×1

$$\hat{y}_i = \sum_{t=1}^T \hat{v}_i \cdot t$$

$$= x_{i,t+1} \hat{\delta}_{OLS} + \frac{\hat{\sigma}_u^2}{\sigma_u^2} \cdot \hat{v}_i$$

$$H_0 \quad \frac{\hat{\sigma}_u^2}{\hat{\sigma}_1^2} \cdot T \cdot \hat{\sigma}_v^2 \equiv \mu_{ii}$$

$$\boxed{\hat{\mu}_{ii} = \frac{\hat{\sigma}_u^2}{\hat{\sigma}_1^2} \cdot T \cdot \hat{\sigma}_v^2}$$

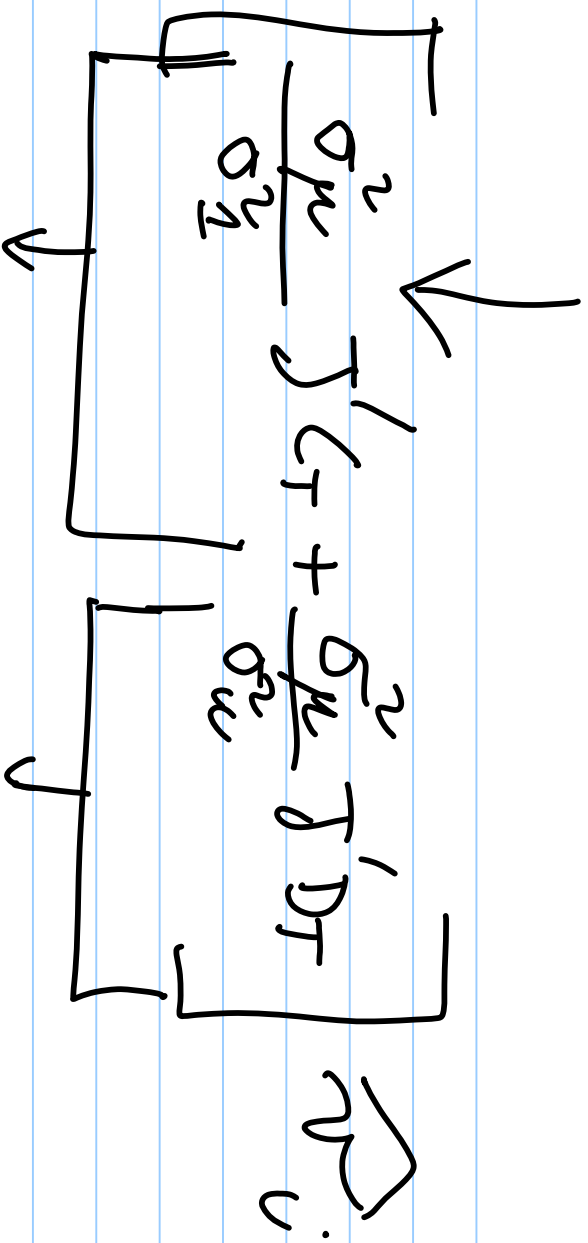
→ STIMULIERUNG
(EFFEKT INDIVIDUELLER PARAMETER)

INTERPRETATIONS OF $\hat{\mu}_c$

$$y_i = x_i \beta + v_i, \quad v_i = \mu_c \delta_i + u_i$$

$$y_{iTH} = x_{iTH} \hat{\delta}_{iTH} + \boxed{Q' V^{-1} \hat{v}_i}$$

price component of v_i .



QUESTA ESSENZIALE EVIDENZIA LA SENSIBILITÀ DEI PARAMETRI μ_i .

NELLA PRATTE INTRINSECA A μ_i E NELLA PRATTE INTRINSECA A U_i .

TESTI DI SIGNIFICANTIA DELA EFFETTI RANDOM (CEVAI)

TEST DI BREUSCH-PAGAN (TRPO LM)

(BP)

$$BP \sim \frac{\sigma}{\sigma_0} \chi^2_{(q)}$$

DAVE

$H_0: \sigma_{\mu}^2 = 0$

(NO EFFETTI RANDOM)

$H_1: \sigma_{\mu}^2 > 0$